

Public and Private Enforcement of Securities Laws: Resource-Based Evidence

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Abstract

Recent work in finance concludes that financial market development is facilitated by private enforcement of investor protection — primarily via disclosure and private lawsuits — but not strongly affected by public enforcement of regulatory rules, or even criminal penalties. Several World Bank pronouncements adopt this view. In this paper, we re-examine the value of public enforcement. In contrast to prior work, which used formal enforcement powers as a proxy for public enforcement, we introduce and use two new measures of public enforcement based on securities regulators' real resources — their staffing levels and budgets. Four standard measures of stock market development (market capitalization, trading volume, the number of domestic firms, and the number of IPOs) are all strongly related to these resource-based measures of public enforcement, with more intense public enforcement regularly correlating with more robust financial outcomes. In horse races between our measures of public enforcement and the most common measures of private enforcement, public enforcement is typically at least as important as private enforcement in explaining important financial market outcomes around the world.

Keywords: investor protection; enforcement; securities regulation

JEL classifications: D21, G14, G18, G24, G28, G32, G34, G38, K22

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Introduction

We evaluate here the value of public enforcement of securities law for the development of stock markets around the world. Properly weighting the relative value of public and private enforcement has major implications both for the academic understanding of what propels strong financial markets and for the content of current development programs. The World Bank, for example, in seeking to promote strong financial markets and economic growth has recently dismissed public enforcement of securities laws as being unimportant, while seeing private enforcement as central, conclusions we consider to have been reached too hastily. World Bank (2006: 1). Djankov et al. (2005) and La Porta et al. (2006) highlight the importance of evaluating the relative importance of private enforcement versus public enforcement of securities obligations that protect investors.

In principle, both enforcement mechanisms could have serious defects and strong advantages. On the defect side: A public enforcement system is degraded because public actors have mixed and often weak incentives to do their jobs well and because they often suffer from poor information of both general market and specific firm conditions. But a private enforcement system is subject to collective action and free rider effects among dispersed investors, to slow and inaccurate judiciaries, to lawyers' rent-seeking, and to the resulting potential inability of private enforcement to visit severe monetary penalties on wrongdoers. On the positive side, public enforcement could be run by public-regarding policymakers and can invoke sharp criminal, financial, and reputational penalties that deter egregious wrongdoing, while private enforcement actions can be brought by actors who are best informed because they are close to the transactions at hand.

There is no obvious a priori winner. Hence, we seek to evaluate whether public enforcement is in practice valuable by using a new data-set of the resources devoted to securities market oversight around the world — resources measured by the staffing levels of securities regulators and by their budgets. Prior

studies relied on indices based on the formal powers of regulatory officials to measure the strength of public enforcement, a measure that may well be less appropriate than money and people.

In Part I, we discuss legal academic analyses on why private lawsuits — a key component of private enforcement that prior work has emphasized — often do not visit penalties on the relevant actors, can distort incentives, and can be inefficacious. Consequently, even suboptimal public enforcement can contribute to good financial outcomes, because the real world's on-the-ground private enforcement is often misdirected. Although public authorities lack ideal incentives and may be more poorly informed than private actors, the common structures of private lawsuits *also* yield poor incentives and weak information for the relevant private actors. Moreover, strong public enforcement enhances the efficacy of private enforcement.

In Parts II and III we introduce our own measures of public enforcement based on regulatory budgets and staffing levels of securities market regulatory and enforcement agencies, and then describe the data and present our findings. First, we explore the relationship between our new measures of public enforcement and the size of a country's capital markets. Allocating more resources is positively associated with measures of robust capital markets, including market capitalization, trading volume, the number of domestic firms, and the number of IPOs. We then control for measures of private law enforcement. Our resource-based public enforcement variables' relationship to stock market development is robust to adding controls for private-law oriented legal factors, such as disclosure requirements and liability rules, which prior financial work has seen to be critical to developing and maintaining strong securities markets. In several instances, public enforcement dominates private measures, such as liability standards, results consistent with current legal thinking that civil litigation, at least in the United States, is poorly designed. In horse races between our resource-based measures of public enforcement and formal measures of public enforcement powers commonly used in prior academic work on the subject, our new measures of public enforcement are more strongly associated with robust financial markets. Moreover, the two prominently-used formal measures of public enforcement do not correlate with one another,

further suggesting that indexing formal regulatory powers may not be the best research methodology to measure real public enforcement intensity.

The data here do not, however, support a conclusion that public enforcement uniformly trumps private enforcement. First, the results soften when we account for influential observations, although when we do so, public enforcement persists as being as, or more, significant than private enforcement. While the liability standards variable associated with civil litigation proves to be largely insignificant, the disclosure variable, which is associated with private enforcement, is, like the resource-based public enforcement variables, often significant. And, in some areas, public enforcement may be less effective than private enforcement, in others the opposite may be true. As an example of the former, our public enforcement variables do not predict country-by-country results of variables associated with highly dispersed ownership of public firms. As examples of the latter, our measures of public enforcement are more highly correlated with a country's market capitalization and its ratio of domestic firms to population.

We also discuss the direction of causality. It's possible that more public enforcement produces better outcomes; it's also logically possible that stronger financial outcomes call forth more public enforcement, with higher budgets and deeper staffing levels. In particular, widespread ownership of financial assets in a democracy should make it easier for government officials to spend resources on public enforcement. If influential elements of the polity own financial assets, they will insist on being protected from insiders' machinations. Unpacking causation is complicated, however, and we see the two aspects — market results and budgets — as likely to be simultaneously determined, with causation bidirectional. But we want to ascertain, to the extent one can, that causation does not run solely from markets to budgets. While the finance literature has not yet settled on a suitable method here for this kind of data — the accepted technique is to use instrumental variables, but finding a suitable one has vexed academic finance for a decade — and while our findings are only tentative, we view our findings on directionality as no weaker than those previously used in this area of finance work. (Differences-in-differences and country fixed-effects regressions are not yet viable, due to data limits.)

In Part IV, we discuss the channels through which public enforcement could improve financial markets and offer cautionary notes for policymakers. Public enforcement strength can indicate the degree to which the public authority can (1) conduct market surveillance, by addressing general systemic issues and problems that need rule-making fixes, (2) investigate individual firms for wrongdoing, (3) bring enforcement actions (including those that result in large financial penalties or criminal imprisonment), and (4) write, revise, and enforce better, more sophisticated regulatory rules. We cannot distinguish which among these four channels for public enforcement is the most important.

Moreover, public enforcement has multiple targets. Much analysis focuses on corporate-governance-related enforcement, especially that connected with tunneling value out from the firm into controllers' hands. Tunneling and related party transactions are important and deserve the attention they have received. But deep public markets also require brokers and others securities-handling institutions that are often intensively regulated in the United States and other countries with deep capital markets. Trading channels have not been the focus of analysis in the law and finance writing of the past decade, yet reliable trading channels might be critically important to building a strong securities market.

* * *

La Porta et al. (2006) and Djankov et al. (2005) are the academic articles closest to our work here, although they conclude that public enforcement is of limited value and we, using resource-based evidence, conclude otherwise. Key policy analysts quickly accepted the relative value of private over public enforcement, with the World Bank advising “[i]n banking and securities markets, characteristics related to private monitoring and enforcement drive development more than public enforcement measures.” World Bank (2006: 1). Researchers now associated with the IMF and the European Central Bank conclude similarly. Bruno and Claessens (2008); Hartmann et al. (2007). Policymakers with such views would promote private enforcement rules and institutions to the detriment of public enforcement institutions.

I. Public and Private Enforcement in Prior Legal Analysis

We were at the outset skeptical of a claimed superiority of private enforcement over public enforcement. Prior financial analyses' faith in private litigation seemed misplaced, as the conventional legal academic view, which we share, is that securities litigation, at least as practiced within the United States, is seriously compromised. Private securities lawsuits in the United States (1) often provide meager returns to wronged plaintiffs, (2) usually do not visit their costs on the wrongdoing actors inside public firms, because the wrongdoers can usually transfer the costs to others, and (3) often just transfer losses from one innocent group of shareholders to another innocent group, with large fees obtained by the lawyers for both sides. Coffee (2006); Grundfest (1994); Romano (2005); Cox et al. (2003). These failings of private litigation reflect the familiar limits on dispersed shareholders' oversight, whether via private litigation, the exercise of voting rights, or otherwise in controlling insider, managerial misbehavior: dispersed ownership creates collective action problems that dilute shareholders' capacity to litigate effectively, to vote efficiently, or to otherwise take remedial actions to control insider misbehavior. Berle and Means (1932); Jensen and Meckling (1976); Jensen (1986, 1989). Private securities litigation via class actions to remedy the shareholder dispersion that disables shareholders from effectively deterring managerial misbehavior often results in class action attorneys controlling the litigation in ways that do not fully benefit dispersed shareholders. More generally, private enforcement cannot provide a wholly or even largely self-sufficient system of securities regulation, as the transactions costs of private remedies for even basic issues like insider trading stymie private enforcement from being fully effective. Litigation-based forms of private enforcement can cover only certain kinds of corporate wrongdoing. Indeed, the major securities reform statutes of the 1930's were enacted to address these shortcomings of private enforcement, Seligman (2003), and the Securities and Exchange Commission has played a major role in American securities markets ever since.

As a matter of casual observation, most well-developed capital markets use multiple public regulatory functions that dispersed shareholders and other private parties are poorly situated to undertake. Detecting insider trading and market manipulation, for example, requires centralized, continuous and

comprehensive oversight of trading markets, functions that neither a single shareholder nor even all the shareholders of a single corporation could maintain efficiently. Disclosure requirements and consistent accounting rules are public goods that private parties on their own would have difficulty devising, updating, and enforcing. Periodic examinations and inspections of broker-dealers are typically the province of government or quasi-public self-regulatory agencies. Systemic risks and liquidity crises are not readily remedied by private contracts or ex post litigation, as the recent Bear, Stearns intervention illustrates, but are tasks for the public regulators. Finally, private parties cannot impose critical sanctions, such as revoking licenses or imposing criminal penalties. While the problems of public administration are familiar and profound – public officials have imperfect incentives and are often poorly informed on many market matters – the widespread use of public enforcement suggests it may still help to develop or maintain robust capital markets. Public enforcement is highly imperfect, but because private enforcement is compromised by free-rider and other weaknesses, public enforcement may still play a useful role.

Furthermore, central elements of private enforcement depend on public enforcement. Consider the centrality of private enforcement based on price movements of securities in response to good and bad news about corporate activities (and the allocative and corporate governance actions that often result). Prompt publication of such news depends on the reliability of the corporate disclosure framework. But in all developed capital markets, public authorities such as the SEC heavily regulate corporate disclosure, and penalize egregious failures of accurate disclosure. And, for shareholders to effectively exercise voting power, another key source of private enforcement, they also typically depend on publicly-enforced, mandatory disclosure. Lastly, class action securities litigation in the United States usually results from the failure of corporations or their officers to comply with publicly-articulated disclosure standards. Recent empirical work suggests that initiating an SEC enforcement action substantially increases the likelihood that private lawsuits will be brought. Cox et al. (2003). Thus, on many dimensions extensive public enforcement appears to be a prerequisite to effective private enforcement.

A final source of our concern with prior work on the subject was the formal manner in which public enforcement was modeled. La Porta et al. (2006) used an index of public enforcement based on

financial regulators' formal characteristics: the financial supervisor's independence from the executive, the supervisor's level of investigative powers, its capacity to issue remedial orders, and the range of criminal sanctions available. They then add these values up to reach an index of what they measure as public enforcement (the LLS index). Similarly, Djankov et al. (2005) developed a second index of public enforcement based on whether the regulator can sanction a specified insider transaction via "(1) [a] fine for the approving body; (2) jail sentences for the approving body; (3) fines for [principal wrongdoer]; and (4) [a] jail sentence for" the principal wrongdoer.

While examining the regulators' formal qualities is a plausible place to begin, it falls short of reliably indicating public enforcement's bottom-line efficacy. One reason here is that familiar problems afflict developing accurate cross-country indices of legal rules, in that it's difficult to obtain consistent classifications across jurisdictions and weight multiple factors appropriately. Armour et al. (2008). Interpretive considerations are also in play, in that the direction of some legal factors included in earlier public enforcement indices seems ambiguous. The LLS index, for example, give more points to an independent regulator, like the American SEC, and fewer to one integrated into a consolidated financial supervisor, like the British Financial Services Authority. While regulatory specialization, such as the SEC's, may be useful, the wider regulatory vision of a consolidated regulatory operation could actually improve supervisory oversight. Prior literature is divided regarding the relative merits of consolidated supervision. Čihák & Podpiera (2006). Hence, it is not obvious that independence trumps consolidation and so doubt arises as to whether the LLS index coding is justified.

A more fundamental problem with an authority-based enforcement index is that it relies heavily — too heavily, in our view — on the regulators' formal legal powers to investigate and sanction. Consider, for example, the power of a securities regulator to refer securities law violations to criminal justice authorities for prosecution. While it plausibly indicates the intensity of public law enforcement, recent studies show that many regulatory authorities have these powers but do not use them, as there are few or no criminal prosecution of securities cases. See, e.g., Jackson (2006). As recent research into the efficacy of insider trading regulation has demonstrated (Bhattacharya & Daouk (2006)), the key issue is

not whether a regulatory body has the *formal* power to sanction offending parties, but whether that power has *actually* been exercised.

For all of these reasons, we set out to explore whether we could devise a more accurate measure of the intensity of public enforcement and then use that measure to consider the question of whether public enforcement had a positive relationship to the development of robust capital markets.

II. Data Description: Measuring Public Enforcement via Regulators' Staffing and Budgetary Resources

We develop in this article several measures of the intensity of public enforcement of securities regulation based on the regulators' budgetary resources and staffing levels. Our resource-based concept of public enforcement efficacy measures the level of public resources a nation allocates to its financial regulators, normalized to reflect either the economic size of the nation or its population. Higher budget and greater staffing allow the regulator to examine allegations of wrongdoing, to write its rules carefully, to conduct market surveillance and review filings, and to act more often to remedy, prevent, and punish wrongdoing. In contrast to prior work, we do not factor in the regulatory agencies' formal degrees of independence from other governmental authorities or on their formal levels of authority to fine or incarcerate wrongdoers. Regulatory independence and high levels of agency authority are of little value to effective enforcement if the agency's budget is minuscule and its staffing thin. And conversely, a not-very-independent regulator with a high budget and strong staffing indicates that political and market authorities have given the agency the go-ahead to enforce financial rules. Similarly, a well-staffed and well-funded agency can, even if it has only limited formal sanctioning authority, make good use of the sanctions that it has.

To be sure, our measures based on budgets and staffing also have their limits, both in concept and construction. Simply because budgetary authorities have allocated ample resources to the regulatory authority does not mean that the regulator deploys these resources wisely or that it even deploys them for the purposes for which they were appropriated. Simply because a securities regulator has ample resources does not guarantee that it utilizes them to bring enforcement actions, to write good rules, and to hire good

people. Deep staffing of sinecures or crony-oriented appointees may lead to a regulator showing a high budget and high staffing levels, but would result in little enforcing of securities regulation. Indeed, we see examples of jurisdictions maintaining relatively similar regulatory staffing levels and budgets, but imposing quite different levels of public sanctions. Jackson (2007) shows that the U.S. and the U.K., with similar resource levels use the resources differently, with the U.S. using more public sanctions and the U.K. focusing on other aspects of supervision. Moreover, lean staffing and a low budget, if the regulator knows how to pick battles and impose severe penalties, can yield strong public enforcement, such that private actors calculate that the probability of being caught times the penalty imposed exceeds the private benefits sought. Thus, staffing levels and budgets may be a noisy proxy for effective public oversight of capital markets. Jackson (2006). But, while they should not be the last word in measuring public enforcement, these variables still improve our understanding here and help to give us a better view on the tension between the legal literature's mixed view of private enforcement and the finance literature's negative view of public enforcement.

We considered shifting to enforcement outcomes, such as indicators of aggregate regulatory actions brought, fines levied, and conviction rates. Staffing and budget are inputs, not outputs. But we did not shift focus for practical and analytic reasons. First, output data is unavailable for many countries. Second, its interpretation would be ambiguous. Low enforcement output could result from the regulators having tremendous budgets and staffing, such that financial actors would not dare break any of the regulator's rules. Third, the mechanisms of enforcement differ across national boundaries — some, like the U.K. and Japan, rely on informal discussion and administrative guidance; others, like the U.S., bring more formal actions against wrongdoers — making variables based on publicly-reported formal sanctions inaccurate measures of regulatory outputs for many jurisdictions. Jackson (2008).

Constructing a budget and staffing data-set comparable across nations is not an easy task. Many jurisdictions divide regulatory responsibility among different bodies, with some bodies operating as self-regulatory organizations and others located within subnational jurisdictions, such as states or provinces. Other nations consolidate capital market regulation in a single financial supervisory agency or in the

nation's central bank. It is often hard to break out the budgets and staffing of securities regulators from those of larger governmental bodies.¹

Notwithstanding these difficulties, we extended Jackson (2007) by assembling resource-based data — budget and staffing — for three samples of jurisdictions. We describe that data in Table 1 and present it in Table 2, drawing primarily from the 2006 edition of *How Countries Supervise their Banks, Insurers and Securities Markets*, which the Central Banking Publications of London compiles annually. Each sample of jurisdictions includes two resource-based measures. One is the number of regulatory staff charged with overseeing capital markets divided by country population; the other is the securities regulatory budget divided by GDP. Our first pair of samples — the most basic ones — consist of jurisdictions for which the *How Countries Supervise* source indicated the regulatory staffing level and budget associated with securities enforcement. This sample yielded staffing estimates for 30 countries and budget observations for 26.

How Countries Supervise (2006) also indicated the budgets and staffing of several consolidated regulators that are responsible for other areas of the financial services industry as well; data for such regulators are not included in the first, basic sample. But for our second sample set, which we denominate our extended samples, we add to the basic sample those jurisdictions where there was objective information with which to allocate staffing and budgets of consolidated agencies to the securities sector, generally through personnel reports included in annual reports or other agency publications. This extended sample includes 39 staffing and 35 budget observations.

Our final sample set, which we denominate the extrapolated samples, adds to the extended samples the remaining jurisdictions for which the Central Banking Publications data included staffing or budget estimates for consolidated agencies, but for which there was no metric to allocate directly the staffing and budget levels for the securities division. For these jurisdictions, we extrapolated staffing and budget estimates based on the median ratio of securities staff or budgets to other areas of financial regulation

¹ Another complexity, which we considered but not incorporate into our analysis, is whether to factor economies of scale into budget and staffing. We normalize staffing levels by national populations and budgets by GDP at current exchange rates.

where direct data on both levels were available.² The extrapolated samples include 53 staffing and 46 budget observations. A description of these and our other variables is in Table 1. We report in Table 2 the six resource-based public enforcement samples, alongside two indices of public enforcement based on the regulators' formal powers that La Porta et al. (2006) and Djankov et al. (2005) develop. Sample construction here is imperfect: The extrapolated sample is the largest, but the noisiest. The extended is more objective, based on data (without extrapolation) from the regulator, but there may be a selection effect in which nations' regulators report such data. The basic set has less noise, without the weakness of the extended's selection bias, but is biased toward nonconsolidated regulators and has fewer observations. Below we indicate some parameters for constructing better measures of regulatory intensity.

We report results using the extended samples as our principal resource-based measure of public enforcement. We do so because the extended samples are the largest samples based on objective, not extrapolated (and, hence, based on less noisy), measures of staffing and budgets. But we also run our principal regressions with the basic and extrapolated samples and include the most important results with these alternative samples in appendices. For all principal findings, the results are consistent across the three pairs of resource-based samples.

Countries allocate sharply differing levels of resources to financial oversight. Some differences reflect the differing roles of financial markets in certain jurisdictions: Financial centers like Hong Kong and Luxembourg have the highest staffing and budgets, and they are jurisdictions where financial activity plays a large role in the domestic economy. But even among more traditional economies, variation is ample. Canada reports nearly 39 regulators staffing their securities agency per million of population whereas Spain, with a comparable GDP, reports only slightly more than 7 staffers per million of population. By this measure Spain's regulatory intensity is a fifth that of Canada.

² As a rough check on this method of extrapolation, we examined the allocation of consolidated staff and budgets (using median ratios) in the extrapolated sample to the extended sample's consolidated regulators' allocations that were based on objective data. In roughly a third of the cases, the allocations were close to the extrapolation ratios; in another third (for countries such as the UK and Luxembourg) the ratios were higher; and in the final third (countries with smaller capital markets such as Iceland and Germany), the allocations were lower. This relatively even distribution of adjustments suggests, at least to us, that our crude extrapolation technique was not obviously biased and, therefore, a useful robustness check on our use of the extended and basic samples.

In Table 3, we report a matrix of pair-wise correlation coefficients for the resource-based enforcement variables and several other variables that we will run, such as indicators of private enforcement and indices from LLS and Djankov of the regulators' formal public enforcement powers. These latter indices seek to measure public enforcement intensity by constructing indices of law-on-the-books, including indicators of the formal powers of the public enforcers. The indicators of regulators' formal powers correlate only modestly with all three sets of resource-based measures — it's generally positive but never greater than 0.5. Accordingly, some countries must be adopting formal rules that could facilitate enforcement, but then fail to allocate the staffing and budget that could make the regulator effective. France illustrates this phenomenon. It has a quite high formal powers index score of 0.77 on the LLS formal enforcement index, more than a standard deviation *above* the mean. But its regulatory budget is only \$29.2 thousand per billion of GDP, less than a third of the mean. If we looked only to the regulator's formal powers and independence, France appears to be *high* on public enforcement, but in our view it should be coded as *low* on public enforcement. This may help explain why our results for public enforcement (as predicting financial strength) differ from prior inquiries. The Netherlands, in contrast, has a slightly below average formal public enforcement authority, but, if we use real resources, it ranks well *above* the mean.

In the correlation matrix in Table 3 we also report the relationship between staffing and budget to other commonly used indices of interest, such as a securities disclosure index, a liability index, an anti-director rights index, and a judicial efficiency index. In general and as would be predicted, our variables are positively correlated with these other indices. Hence, some of the good financial results previously associated with these three formal legal indices could be due to their correlation with the intensity of regulatory enforcement.

One final point of interest in Table 3: The two leading indices of formal enforcement powers do not correlate with one another. Their pairwise correlation coefficient is negative (-0.01). As we discussed above, there are questions about the theoretical value to enforcement of the components with which these public enforcement indices are constructed. The lack of correlation between the two most

commonly used indices of public enforcement in the finance literature raises questions about their robustness for statistical analysis.

III. Results

Our basic research pattern is first to replicate prior analyses of public enforcement from La Porta et al. (2006), substitute our new resource-based measures of public enforcement for the prior enforcement indices, and then check for robustness, including robustness when controlling for measures of private enforcement. The results throughout have resource-based enforcement strongly and significantly correlating with the size of domestic capital markets. The resource-based enforcement variables are consistently as strongly associated with robust capital markets as the best performing index of private enforcement (disclosure) and substantially more strongly associated with robust capital markets than several other indices of private enforcement, including liability rules and anti-director rights.

A. Replicating Prior Work, but with Resource-Based Public Enforcement Measures

In Table 4, we replicate and then re-examine prior work that sees public enforcement as not having much impact on financial market development, see La Porta et al.'s (2006), following them in first using their primary measures of capital market strength: stock market capitalization, trading volume, number of domestic firms, and the number of IPOs.

Within each panel, we run 5 sets of regressions for each capital markets measure. In each set, we first replicate the original LLS regression (column 1) on enforcement, then substitute our resource-based variable for the LLS formal enforcement index (columns 2 and 4), and then present an additional regression with both a real resource variable and a formal public enforcement index (columns 3 and 5). In columns 2 and 3, we employ the staffing variable from our extended sample; in columns 4 and 5, we use our budget variable from the extended sample. We repeat this set of regressions for each of the 4 variables measuring the robustness of capital markets.

Overall, resource-based enforcement (significant in 14 of 16 appearances in Table 4) is superior to the LLS formal enforcement index (significant in only 4 of 12 appearances) in predicting the 4 measures

of robust capital markets used in this analysis. In the 8 instances in Table 4 where both measures of public enforcement appear (columns 3 and 5), our resource-based measures are statistically significant 6 times; formal enforcement is significant only twice. When we run our resource-based variables without the formal public enforcement variable, they had statistically significant coefficients in all 8 cases; the formal index had statistically significant coefficients in 2 of its 4 solo runs.³ In short, resource-based enforcement dominates formal enforcements indices in these regressions.

Introducing resource-based public enforcement affects the prior variables' significance. Consider Panel D of Table 4, where the dependant variable measure the level of IPOs. In the original formulation with formal enforcement index, that index was significant at the 5 percent level. When we add either resource-based enforcement variable to the regression, resources were significant, but the coefficients for formal public enforcement index lost significance. In Panel C, measuring the number a domestic firms, the anti-directors index, a measure of private enforcement, similarly lost statistical power in the face resource-based public enforcement. Hence, as the resource-based measures of enforcement often moderately correlate with these other variable (see Table 3), the absence of real resource variables from earlier work may have produced an omitted variable bias.

B. Public Enforcement, Controlling for Private Enforcement

We next see whether resource-based public enforcement is robust to private enforcement. We run regressions similar to those in Table 4, but add the two private enforcement indices that have been seen as most strongly associated with robust capital markets: disclosure and liabilities standards. We use as dependent variables more current 2004 data from the World Bank databases, because we are no longer seeking to replicate the prior LLS work on formal public enforcement (which used more dated measures of capital market development).⁴ (We use a contemporary wealth control adjusted for purchasing power

³ In Tables 5 and 6, we run similar tests using updated data in a more demanding fashion by adding private enforcement measures. Results for resource-based enforcement are similar to those in Table 4.

⁴ These dependent variables are analogous substantively to those used in La Porta et al. (2006), whose results we sought to replicate and extend in Table 4 but are more recent, based on 2004 data rather than averages for their 1996 to 2000 data. Because several of the underlying independent variables are now more recent measures, we used contemporary wealth controls whenever possible. Also, our domestic firms' measure is based on the ratio of domestic firms to population, not the log of that ratio.

parity,⁵ a control for judicial efficiency, and a control for corporate law with Djankov et al.'s (2005) antidirector rights index.) Results are in Tables 5 and 6.

Prior analysis in Djankov et al. (2005) and La Porta et al. (2006) showed an association between disclosure and liability standards on the one hand and robust capital markets on the other. We replicate these findings in the first four columns of Panel A in Table 5. Here, disclosure displays a consistently statistically significant relationship to market capitalization. But public enforcement, measured by the index of formal powers, does not. The results in the next five columns of Panel A, however, call into question the robustness of those replicated findings. Once we add our real-resource variable for staffing in columns 7 and 9, the coefficient for disclosure remains positive, but loses its significance. And, in columns 8 and 9, the coefficient for the private liability index is negative in the face of resource-based public enforcement, a result inconsistent with prior writings on the subject in the finance but consistent with the dominant critique of private securities liability rules in the legal literature.⁶ Through all five formulations, our resource-based public enforcement variables retain statistically significant, positive coefficients.

The results are similar in the other panels of Table 5, where we present analogous regressions for our three other measures of robust capital markets: trading, domestic firms, and IPO's. In 17 of 20 cases in Table 5, our resource-based staffing variable has statistically significant coefficients. While the disclosure index almost always stays positive, it often lacks statistical significance in the face of resource-based enforcement. The liability index is rarely statistically significant and regularly turns negative.

The results in Table 5 are consistent with the many alternative formulations we discuss below and note in the Appendix: both our real resources variables and the LLS disclosure index are consistently associated with robust capital markets. Our public enforcement variables outperform the disclosure index

⁵ Although log GDP/capita is common, purchasing parity indices have been used to more accurately show wealth contrasts. The choice of wealth controls does not affect results.

⁶ To be clear here, we find significant positive results elsewhere associating disclosure rules and financial outcomes. And our priors are that disclosure rules (whose efficacy depends in part on public enforcement) are a key part of securities enforcement. Our results for liability rules are overall not supportive of its importance and those results are consistent with our own prior — that liability rules may be important for litigation outcomes, but are not important to securities market depth and breadth.

when the dependent variable measures the ratio of domestic firms to population (Panel C). While the disclosure index sometimes has a stronger relationship to market capital than shown in Panel A of Table 5, real resources also generally perform better in this functional form. The two are roughly equally matched in predicting the value of IPOs to GDP (Panel D). The disclosure index, in contrast, often performs better in predicting trading levels (Panel B).

In Table 6, we run analogous regressions with our extended budget sample. The results are quite similar: resource-based public enforcement is always significant; the results for private enforcement are mixed, with disclosure insignificant in 6 of the 8 relevant runs (turning negative once) and liability only significant once, but then with the wrong sign..

The resource-based results are significant economically as well as statistically. The average ratio of stock market capitalization to GDP in the extended sample is about 83. The coefficient for our budget variable (.77) in Panel A of Table 6 implies that an increase of one standard deviation in the level of an average country's budget would be associated with a doubling of stock market capitalization. The implied impact of an increase of one standard deviation in the securities disclosure index would be to increase an average country's market capitalization to GDP by less than 10 percent.

C. Robustness Checks and Tests With Alternative Samples

These results look encouraging for the view that public enforcement intensity has been implemented sufficiently satisfactorily to play a role in keeping financial markets strong. We next examine whether these results hold up over the two alternate sample specifications: the basic one and the extrapolated one. The results are robust, with the basic sample's resource-based variable doing better than the extended sample's and the extrapolated sample's doing somewhat worse.

1. The Basic and Extrapolated Samples. For the basic staffing sample of 30 observations, the dataset typically had sufficient data for the other variables for 28 observations to be present in three panels, but only 20 in the fourth, due to limitations in the number of observations for IPO values to GDP. As before, we ran the public enforcement as an independent variable in 4 panels, each with one of the 4

indicators of financial market strength as the dependent variable. Each panel had 5 separate specifications: resource-based enforcement alone (with just the controls), with formal public enforcement, with disclosure, with liability, and then against both disclosure and liability. The basic staffing variable was significant in all 20 specifications. Disclosure was significant in only 3 of the 8 specifications in which it appeared. Liability did poorly: always negative, significantly so in 4 of the 8 specifications in which it appeared. These results are presented in the (unpublished) Appendix, in Table 5A. Figures 1 and 2 present residual plots, illustrating the significant relationship between national financial outcomes and basic resource-based public enforcement, after controlling for private enforcement, anti-director rights, judicial efficiency, and wealth, in the form of the model in column 9 of Table 5. (For completeness, Figures 3 and 4 present residual plots for the extended resource-based sample and Figures 5 and 6 do so for the extrapolated sample.)

For the extrapolated staffing sample of 53 observations, 46 were present in three panels, 36 in one. The extrapolated staffing sample, conceptually the noisiest of our samples, was significant in 10 of the 20 models, negative in one. Disclosure was significant in 4 of the 8 specifications in which it appears, negative in one. Liability again did quite poorly, significant only twice, but with the wrong sign. These results are in Table 5B of the (unpublished) Appendix.

For the basic budget sample of 26 observations, 24 were present in three panels, but only 17 in one. Resource-based public enforcement, as measured by the basic budget data was significant in all 20 specifications, disclosure significant in 4 out of its 8 appearances. Liability turned negative in 4 out of its 8 appearances, significantly so twice. These results are in Table 6A of the (unpublished) Appendix.⁷

Lastly on this robustness check, we reran our regressions on the extrapolated budget sample of 46 observations. In 3 of the 4 panels for outcomes, 39 observations survived, 31 in the fourth panel. The extrapolated budget variable was again significant in all 20 specifications. Disclosure was significant twice in its 8 appearances, turning negative once. Liability was again erratic, turning negative in 4 of its 8

⁷ We also normalized by population (instead of GDP), because 3 of the 4 primary outcomes had GDP denominators. The results, reported in unpublished Appendix Table 6D, were substantially similar.

appearances, significantly so once. These results are in Table 6B of the (unpublished) Appendix.

Overall, the resource-based measures of public enforcement are robust in predicated financial market size and depth across three sample specifications for each of two resource measures.

2. Influential observations. With the relevant samples no larger than 53 observations, we examined the potential impact of influential observations. Hong Kong, Luxembourg, and Singapore have deep and broad financial markets and devote significant resources to supporting their markets. At one level, these observations are consistent with resource-based public enforcement playing a key role in financial markets. Hong Kong, for example, has the reputation of having one of the toughest public securities regulators, one that regularly applies tests of substantive suitability (a stronger standard than that of adequate disclosure). But we test for the effects of their influence nevertheless, as these small nations and their role as financial centers may in some ways differ from what happens in the rest of the world. The figures show Hong Kong as an influential point, but a visual relationship persists between the other nations' financial outcomes and enforcement resources. To more formally test whether their influence drives the results, we first re-ran the extended sample results using STATA's robust regression, which systematically clips or eliminates outliers.

The robust regression results for the extended staffing sample are in Table 5C of the (unpublished) Appendix. The extended staffing variable is significant in 9 of the 20 specifications but turns negative in several specifications with trading volume as the outcome. Disclosure is significant in 6 of its 8 appearances. Liability again is erratic, with a negative coefficient in 3 (significantly so once) of its 8 appearances. For the extended budget sample, the results (in Appendix Table 6C) are similar (although without any negative coefficients on resource-based enforcement for trading volume, or elsewhere and with the disclosure index having only 3 of 8 formulations with statistically significant coefficients).

As a further robust check, we ran Cook's D for influential points and outliers. The test often identified Hong Kong and Nigeria as influential observations. Conventional econometric technique is to examine these observations for accuracy, to consider eliminating them if inaccurate, and to check for robustness with a test such as STATA's `rreg` command (or a similar clipping technique, such as

winsorizing, to preserve the basic information but dampens its impact). Belsley, Kuh, and Welsch (1980: 3); Hogg (1979: 114); Kennedy (1998).. The Hong Kong observation seems accurate and consistent with reports from international securities lawyers we consulted who report Hong Kong to have one of the most intense securities regulators in the world, using regulatory techniques (such as merit review) not used even in the United States.

The Nigerian observation, though, looks suspect. It shows a budget/GDP number more than 4 times larger than that of the United States, an order of magnitude greater than countries with similar wealth. A persistently misplaced decimal point is not impossible. Nevertheless, we looked at the results of temporarily dropping Nigeria first and Hong Kong (and the two other financial centers with strong resource-based enforcement, Luxembourg and Singapore) second. In the latter, the resource-based results do weaken, as does the combined impact of the private enforcement measures (disclosure and liability). In the econometrically favored technique for dealing with accurately-measured influential observations (STATA's `rreg` or similar), the resource-based and disclosure coefficients do well, the liability ones do not, as Table 7 shows. Without the Nigerian outlier, whose accuracy is suspect, resource-based enforcement's significance approaches 100%, disclosure declines in significance, and liability (which was rarely significant to begin with) remains insignificant. Results are reported in Appendix Table 7A-B.

We also winsorized the relevant variables and ran `rreg` with the basic and extrapolated samples. We summarize the effects on the coefficients for all three samples in Table 7. Winsorizing and `rreg` reduce the frequency of significant outcomes for resource-based public enforcement, but about half remain significant, and often the loss of significance is at the .15 level for the p-value. Disclosure does not suffer from a similar decline, but the overall distribution of coefficients for the disclosure index is roughly comparable to the distribution of resource-based coefficients when winsorizing and using STATA's `rreg` regressions. Liability, the other private enforcement variable, again performs erratically. Overall, both resource-based public enforcement variables and the disclosure-based private enforcement do well; the private enforcement liability variable does poorly.

D. Financial Variables Associated With Dispersed Ownership

So far, we have limited our analysis to dependent variables associated with the size of a country's capital markets: stock market capitalization, turnover, number of domestic firms, and IPOs. All of these indicate an important role for public enforcement, as measured by the real resources of staffing and budgets as opposed to formal authority. Other outcome variables measure financial depth and we turn to those now, focusing on three dependent variables closely related to the dispersion of ownership: an index measuring ease of access to public markets, median block premiums, and ownership concentration.

Some analysts have seen dispersed ownership (and low ownership concentration) as closely linked with strong capital markets. La Porta et al. (2006). Hence, they seek to measure capital markets' robustness by the degree of dispersion (and the limited level of ownership concentration). This perspective is contested, however, in the legal literature and some of the financial literature, because concentrated ownership can reduce agency problems of corporate managers and, when potential agency costs are systematically high, one should expect to see more concentrated owners, even if enforcement is effective. See, e.g., Jensen (1986, 1989), Gilson (2006), Roe (2002).⁸

In Table 8, we regress these three variables associated with dispersed ownership on our two key measures of public enforcement and several other legal indices and controls. Resource-based public enforcement does not significantly correlate with the variables associated with dispersed ownership;⁹ the disclosure variable does. While these results require further study, at a minimum, one might conclude that to the extent that public enforcement positively influences the size of a country's capital markets, it does not directly disperse share ownership. Countries that dedicate more resources to securities enforcement do not necessarily have fewer closely held firms or lower premiums for control, although they have larger stock markets, more IPOs, and more firms overall.

⁸ Since the correlation between block premiums and ownership concentration is only about .5, the data is consistent with some nations having low block premiums but modest ownership separation.

⁹ The ownership concentration measure indicates the number of firms in the sample that have a shareholder with more than 20% of the company's stock. Ownership is viewed as dispersed if no such owner exists.

Several possibilities may be in play here: One is that some nations with substantial concentrated ownership need more public enforcement resources to better control insider behavior, thereby generating a negative relationship between regulatory intensity and dispersed ownership. Another is that public enforcement is most efficacious in dealing with the institutions of securities trading — brokers and dealers — and less effective than private remedies in dealing with corporate governance issues. Another is that the method of aggregating ownership concentration in country averages is econometrically incorrect; Holderness (2008) criticizes the aggregation methods prevailing in the literature.

E. Limits to Both Private and Public Enforcement: Intermediate Financial Variables

We also ran these regressions using as dependent variables the World Bank measures of capital market development, involving the size of equity markets, ease of access to equity markets, stock market efficiency, and stock market stability. The World Bank data also has an aggregate index that includes all information about a country's capital markets. We report the regression results in Table 9. In each regression, we included key enforcement variables, the private-remedy-oriented disclosure index, and a wealth control. The results are largely consistent with our previous findings.

In terms of the market size index and the World Bank's aggregate equity index, both our resource-based public enforcement variables and the disclosure index have positive and statistically significant effects. So, as we report above, it seems that both public enforcement and disclosure are associated with more robust capital markets. For the Equity Market Access Index — reflecting the ease with which companies can raise capital — the disclosure index has a positive and statistically significant coefficient, but our enforcement variables do not. Yet *neither* the public *nor* the private enforcement variables correlate closely with market efficiency and stability, as we report in Table 9.¹⁰

¹⁰ We sought to see whether strong relationships existed with components of these World Bank efficiency indices, such as trading costs or price synchronicity, and either legal origin or public enforcement. We found no strong relationship between these on the one hand and either legal origin or public enforcement on the other. Nor did we find strong effects on national costs of capital for either our public enforcement variables or the private legal indices. Other recent work on this subject found several of the LLS legal indices, including their public enforcement index (of formal powers), to be associated with lower capital costs. See Hail and Leuz (2006). Other research found a relationship between some of the LLS legal indices and lower trading costs. Eleswarapu and Benkataraman (2005).

Since the regressions overall show both public enforcement and disclosure to be associated with the size of a nation's capital market, one might expect that either or both would correlate with improved technical performance of the securities markets. Perhaps, because our variables measure money and staff, but not overall effectiveness, a better-tuned public enforcement variable would yield different results. Yet the private-enforcement variables fail to predict market efficiency and stability any better. Since neither enforcement variable does well here, we see this non-result as another reason to be cautious in rejecting public for private enforcement in policy-making.

F. Legal Origin and Regulatory Intensity

The relationship between legal origin and robust securities markets has been of considerable interest in recent scholarship, see, e.g., La Porta et al. (1998), Roe (2006). As Jackson (2007) and Roe (2006) observe, the level of public resources devoted to financial regulation is generally higher in common law than in civil law countries, although only weakly so. This phenomenon is surprising in that the stereotypical view of civil law countries is that they regulate their economies more heavily. Segal and Whinston (2006: 1) explain:

There are two basic approaches to deterring socially harmful behavior: with the threat of litigation by private parties or with enforcement by public agencies. Both approaches are used in most countries, but in varying degrees. Private litigation is common in the United States and (to a lesser extent) the United Kingdom and other "common law" jurisdictions. In contrast, the "civil law" countries, such as those of continental Europe, have far less private litigation, and rely more on enforcement by public agencies.

For financial market regulation and litigation, the opposite occurs, as measured by resources allocated to the regulatory task. Resource intensity, when controlling for wealth, roughly divides along common law/non-common-law grounds, significantly so in half of the six resource-based samples. As Table 10 indicates, common law jurisdictions expend more effort on, and devote more resources to, public enforcement, challenging traditional assumptions about the primacy of public, state power in civil law jurisdictions. Hence, key institutions that seem to support financial markets are not those traditionally associated with common law's strengths.

In Table 11, we explore whether common law also positively correlates with robust capital markets once one controls for real resources expended on public enforcement and the standard private law indices. We re-ran our principal regressions with full controls (column 9 in Table 5 for staffing and column 5 in Table 6 for budget), using our four key dependent variables with both staffing and budget variables from our three samples, yielding 24 separate regressions. We summarize the coefficients on common law for these 24 regressions in Table 11: common law is rarely significant and is often negative, sometimes significantly so. The significance of the coefficients on resource-based public enforcement and disclosure was substantially unchanged.¹¹

Table 11 thus suggests that a country seeking securities market development can achieve that effect by adopting effective disclosure-based *and* real-resource public enforcement regimes that support capital markets. Public and private enforcement, strategies traditionally associated respectively with civil-law-style mechanisms on the one hand and common-law-style mechanisms on the other, both seem to play roles. Once one controls for a limited number of plausible legal mechanisms to reduce corporate and financial wrongdoing, the relationship between common law origins and capital market development fades away. This resonates with recent important analysis in the legal literature, to the effect that the set of basic corporate problems is small, as is the number of institutional solutions. Kraakman et al. (2004).

G. The Direction of Causality?

As is usual in these finance inquiries, we cannot reject the possibility of reverse or bidirectional causality. In fact, our prior here is that causation is bidirectional, with strong financial markets inducing governments to protect an important constituency and an important market sector. Strong financial markets may emerge for reasons exogenous to the intensity of public enforcement (such as economic conditions, private ordering, strong property rights, or the absence of strong alternative financial channels) and then the relevant players (such as founders, investors, and public servants) call forth strong budgeting and more staffing in financial market regulators to protect these already-developing financial

¹¹ Running the same specification using STATA's `rreg` produces analogous results, as we report in Table 11A of the Appendix.

markets with stronger regulators. But we want to see if we can reject the possibility that causation only runs from market breadth to enforcement strength. Persistent correlation alone does not allow us to do so.

The finance literature in the past decade relied on legal origin to demonstrate causality as running primarily from a legal variable to the financial outcome. The theory was that legal origin, having preceded current financial outcomes by centuries, could not have resulted from current financial markets. With legal origin often correlating with the legal variable being studied, especially variables relating to private litigation, the relationship makes legal origins a potentially useful instrumental variable, either formally (La Porta et al. (1998)) or informally (Djankov et al. (2005)). Since the theory was that private enforcement mechanisms associated directly with the common law (such as fiduciary duties) or indirectly with it were driving financial results, there was some plausibility to using legal origin as an instrument.

But problems afflict using legal origin to anchor causality here. While staffing and budget correlate with legal origin, the causal connection between common law origin and high government spending and deeply-staffed regulators is more obscure than for other legal-financial connections based on, say, common law fiduciary duties. The causality of corporate law's effects on finance through fiduciary duties *could* be plausibly tied to legal origin, because common law systems traditionally used fiduciary duties and civil law systems are thought not to use anything similar. Similarly, causality could be associated with the disclosure variable, since disclosure could be seen as arising out of common law fiduciary concepts. And the causality of *heavy*-regulation regimes and poor financial outcomes could be tied to legal origin, because civil law systems are frequently seen to regulate their economies more extensively than do common law systems. But it's the *common* law systems that budget more for regulation and hire more regulators, making the legal origin nexus faint. While it's possible that the underlying idea is that common law systems just prefer financial markets, this seems too attenuated to serve as an instrumental variable anchor. After all, we are no longer talking of institutions — like fiduciary duties, judicial enforcement, and a common law aversion to regulatory intensity — but underlying preferences. But preferences are only weakly tied to legal origin, since the legal origins originated long *before* financial markets were well developed, and many intervening events surely affected preferences. Perhaps because

of such concerns, even proponents of using legal origin as an instrument seem to have withdrawn their support here. Djankov, Hart, McLeish, and Shleifer (2006: 35) (“Since legal origin influences many of the independent variables in these specifications, we cannot properly run instrumental variable regressions, and hence cannot claim that these effects are causal.”).

Nevertheless, because of legal origins’ prominence in the finance literature and prior readers’ comments, we first ran two-stage least squares regressions using legal origins as an instrument, using the full model of the forms of column 9 in Table 5 and column 5 in Table 6. *Neither* public nor private enforcement indices such as disclosure are effectively instrumented by common law in Appendix Table 11A. The coefficients on the variables of interest were never significant and often negative. This lack of significance, and the frequency of the unexpected sign, occurs for both of the resource-based public, and the disclosure- and liability-based private, enforcement measures when instrumented by origin. Hence, in our full models, the common law instrument failed to produce evidence of causality for either public or private enforcement, a result consistent with recent doubts expressed by some of origins’ original proponents (Djankov, Hart, McLeish, and Shleifer (2006: 35)) about origins’ usefulness as an instrument.

We also investigated a second instrument: the log of GDP. As Jackson (2006) reports and as we confirm in Table 12, there are economies of scale in the regulatory staffing and regulatory budgets: The larger a nation’s economy, a smaller regulatory staff or less regulatory expenditure could do the job as well. This phenomenon is reflected in a consistently negative correlation between the logarithm of national GDP and resource-based public enforcement and makes it a potential instrument to explore causality. Scale effects appear to exist for some private enforcement indices, such as the LLS disclosure index utilized in many of our previously reported regressions: the larger a country, the more extensive its disclosure requirements; hence, log of GDP is a plausible instrument for private enforcement as well.

Table 12 reports the results of our instrumental variable regressions using two-stage least squared regressions for the four dependent variables used above to measure robust capital markets: market capitalization, trading, domestic firms, and IPOs. The instrument is log of GDP and we include the full set of controls as before. In Panel A, we report a series of second-stage regressions along with the

original OLS regressions (columns 1 and 3) and then the coefficients from second-stage regressions using first extended staff (column 2) and then extended budget (column 4). In Panel B, we report the corresponding first stage regressions.

While one must be careful about making strong inferences from analyses of the sort presented in Table 12 (see Murray (2006)), the results suggest that our new real resource enforcement variables play a causative role in the development of robust capital markets, at least for the dependent variables other than trading. The most straightforward evidence of the relationship is the number of statistically significant coefficients associated with our enforcement variables in 6 of 8t cases. The strength of our inferences here is tempered by the number of regressions that have Cragg-Donald F-statistics beneath the relevant Stock-Yogo critical values, suggesting that the instrument is weak. However, both the statistical significance of the coefficients on many of our real resources variables and their consistency of these estimated coefficients with the coefficients in the original OLS regressions increase our confidence in the usefulness of this analysis, as Murray (2006) suggests such results should.

We hardly claim to decisively resolve the question of causality. The log of GDP instrument is not ideal, and so the evidence of a causal link between the new enforcement variables and robust capital markets is only suggestive. However, the evidence in Table 12 is as strong as the similar instrumental variable analysis that has been offered to support the causal link between private enforcement and robust capital markets. More anecdotally, regulatory budgets rise after market problems — scandals and crashes — and shrink during market highs. The American experience during the 1990s illustrates: When the stock market did well, Congress cut the SEC's funding. After the Enron and related scandals indicated weaknesses, Congress raised the budget and expanded the SEC's staffing. Increased public enforcement resources did not follow strong market results, but followed financial market weaknesses. We understand this to be a familiar pattern around the world. Moreover, the securities markets that have been most successful in attracting new listings in recent decades — the United Kingdom, Hong Kong, Luxembourg, and, until recently, the United States — have relatively high levels of real resources dedicated to securities enforcement. To the extent that the burden of our argument here is that academic writers and

policymakers should not yet dismiss public enforcement as a potentially central source of capital market development, our instrumental variable analysis buttresses that claim.

IV. Discussion

Here we discuss the potential importance of public enforcement and why prior research led to contrasting conclusions.

A. The Importance of Public Enforcement

Although our method is in the spirit of the cross-country legal methodology that La Porta et al. (1998) pioneered in finance, the public enforcement results we report above differ from La Porta et al.'s (2006: 7-9), who report the following principal findings:

Perhaps most interestingly, both disclosure requirements and liability standards are positively correlated with larger stock markets. ...

The results for public enforcement ... are less consistent. Public enforcement only matters for the external-market-capitalization-to-GDP ratio and IPOs, although it has a large economic effect on both variables. ... In contrast, anti-director rights, but not public enforcement, matter for the number of firms, block premium, and ownership concentration.

These results suggest a preliminary view of what works, and what does not, in securities laws. *Public enforcement plays a modest role at best in the development of stock markets. In contrast, the development of stock markets is strongly associated with extensive disclosure requirements and a relatively low burden of proof on investors seeking to recover damages resulting from omissions of material information from the prospectus.*

La Porta et al. (2006: 19-20) (emphasis supplied). The authors conclude: "*All the evidence*," they state, "suggests that relying on public enforcement is unlikely to be a useful strategy for jump-starting the development of securities markets in poor countries." *Id.* at 25 (emphasis supplied).¹²

The results here with resource-based measures of public enforcement sharply differ. The private enforcement liability standards are regularly insignificant and regularly with the sign reversed (sometimes significantly so). Resource-based public enforcement is regularly associated with deeper securities markets, as strongly as is disclosure. There's no significant evidence here that liability standards play a role in developing financial markets. And, public enforcement appears to have a strong and significant

¹² Their reference to "all the evidence" is puzzling as their paper includes several results that show even the formal-powers-based public enforcement having a positive impact: a strong correlation between their measure of public enforcement and both external-market-capitalization-to-GDP ratio and IPOs, with public enforcement having "a large economic effect on both variables." La Porta et al. (2006: 2).

association with securities markets, suggesting that it could well play a role in developing securities markets in poor countries. When measured in terms of real resources, public enforcement has a consistent and robust association with four key dependent variables that measure robust external capital markets. The results persist across three different samples and using both staff as a percentage of population and budget as a percent of GDP as a measure of public enforcement.

B. Channels from Public Enforcement to Financial Outcomes

While our results show that public enforcement is clearly associated with important financial market outcomes, we are less sure of how it does so, because it can affect financial markets through several channels. First, high budgets and staffing make it easier for the public authorities to conduct market surveillance, by looking for wrongdoing and problems that need rule-making fixes. Market actors would presumably understand that authorities would be more likely to detect any deviant and punishable transgressions in nations where budgets and regulatory staffing are higher. Second, once wrongdoing is detected, higher budget and staffing facilitate regulatory investigations, making it easier for the agency to bring enforcement actions, including those that result in large financial penalties and even jail terms. The penalties for these enforcement actions are also the indirect ones from public enforcement damaging the firm's organizational and reputational capital. Recent studies have shown that the reputational damage, as trading partners impose market penalties, is quite high, often exceeding the direct financial penalties at least in the United States. Alexander (1999); Karpoff et al. (2006). Third, a higher budget and more staffing facilitate the regulatory agency being able to write, revise, and enforce better, more sophisticated regulatory rules. The completeness of legal rules has been shown to affect the efficacy of regulation of insider trading (Beny (2005)) and could be important more broadly. Fourth, much public enforcement is done informally, through administrative guidance in Japan or a regulator's raised eyebrow in England. Britain, with an important securities market, relies little on private litigation, Armour, Black, Cheffins, and Nolan (2007), and much on informal public enforcement. Even self-enforcement, such as Britain's well-known City Code on takeovers, comes from a public-private panel that includes representatives from the Bank of England and the stock exchange, as well as private players, and was for years convened by,

and physically met in, the Bank of England. Armour and Skeel (2007). Such informal public enforcement requires highly trained staffers.

Several market channels could also link public enforcement to stronger capital markets. One is the policing of the wrongdoing-potential of traditional, insider corporate governance actors who take value from outside investors unless the enforcement machinery prevents them from so doing. A second is that financial markets also depend on relatively low-cost, transparent trading mechanisms. Regulatory budgets and staffing could make trading less dangerous for the typical shareholder, thereby deepening financial markets. Again, budget and staffing aggregates do not distinguish which market channel is most sensitive to budgets and staffing. The fact that public enforcement is closely associated with the size of capital markets but not the core corporate governance features that support capital markets is suggestive that this — trading regulation as opposed to just corporate governance — may well be an important and under-appreciated enforcement channel. However, as the analysis here did not establish a consistent association between regulatory resources and technical measures of market performance, public enforcement's role in enhancing trading channels is a theoretical possibility lacking empirical validation.

Finally, the precision of regulation may relate to our measure of public enforcement intensity. When we see simple, bright-line rules in a nation, we could attribute this result to several causes. A popular one recently has been legal origin, with civil law nations using bright-line rules and common law nations using more subtle, fiduciary-based ex post decisionmaking. Another is that simple, bright-line rules are the kind that financial regulators with a weak budget and low levels of staffing can write. If there are only a few trained people inside the regulatory agency, and where the budget for rule-making is low, well-intentioned regulators may be forced to use bright-line rules with sporadic enforcement. Where the budgets are stronger and the staffing deeper, the agency can write more nuanced, tailored rules.

C. The Significance of Certain Mechanisms of Private Enforcement, Insignificance of Others

Our findings are consistent with the value of important private enforcement mechanisms (but not to the exclusion of public enforcement). Consistently with the prior findings of La Porta et al. (2006), the

quality of disclosure is associated with strong capital markets.¹³ And in some of their tests – those measuring disperse ownership and the value of block premia – private enforcement often had a statistically important association, whereas public enforcement did not.

But liability in private lawsuits does poorly in the broad range of our analysis. In fact, when the liability index is run against real-resource measures of public enforcement, the liability index was often negatively correlated with robust capital markets, at times significantly. This result, while sharply in conflict with prior financial economic studies, resonates with legal academic analysis casting doubt on the efficiency of U.S. private securities litigation to police corporate misbehavior. (In diffusely owned firms, innocent shareholders often effectively bear the financial burden of such lawsuits, insiders can often shift payment of any of their own liability to the corporation itself, and lawyers can often direct the lawsuits to their own advantage but not to the best advantage of shareholders and financial markets.) Our results on the unimportance of liability standards are, consistent with the fact that private securities litigation is not an important component of capital market oversight in many well developed securities markets, such as those in the United Kingdom. Armour et al. (2007). That disclosure indices are more strongly associated with robust securities markets than liability indices suggests that private enforcement mechanisms associated with ex post litigation may not be as important that private mechanisms associated with stockholder voting and pricing accuracy. These may have implications, which we do not explore here, on institutional design: The World Bank's focuses on ex post private litigation, which has had some academic support, may be over-emphasized.

Our goal here is not to show that the intensity of public enforcement is *more* closely associated with strong capital market than other factors. Rather we seek to show that its intensity predicts the size of a country's capital markets and, as Tables 5 and 6 show, the relationship between resource-based public enforcement and the size of a country's capital market is comparable to that of disclosure-based indicators and stronger than that found for previously-prominent private litigation basics, such as liability standards.

¹³ As we noted above, disclosure itself depends partly on the efficacy of public enforcement in reducing the inaccuracies of private disclosure and in standardizing the comparability of company-by-company disclosure.

While public and private enforcement interact in important ways, they do not interact simply via higher public enforcement resources primarily going into designing better private liability systems. (Even if they did primarily interact in that way, the results here then suggest that the intensity of continuing *public* design work is what would make private enforcement effective.) A basic familiarity with the American Code of Federal Regulations as it relates to securities laws reveals detailed, step-by-step instructions on what firms must do when selling securities. The bulk — more than half — of the American SEC’s budget goes into basic enforcement activity. Regulators examine financial firms, review issuers’ filings with the regulator, and bring enforcement actions against wrong-doers. Banking and insurance industry regulatory intensity is high in nations where securities regulation is intense. Jackson (2007: 271). For both banking and insurance, even in the United States, private liability is trivial. And outside of the United States, securities regulators are rarely involved in judicial, private enforcement because there is very little of that, as Armour, Black, Cheffins, and Nolan (2007) show, so budget and staffing are likely to map onto the intensity of public enforcement. Bhattacharya and Daouk’s (2006) result on the relationship between insider trading enforcement actions and securities markets’ strength is consistent.

Overall, our results reveal a robust relationship between the intensity of public enforcement and the size of a country’s capital markets. This relationship persists even when we control for legal indices measuring disclosure, liability standards, and judicial efficiency. Disclosure persists as often significant; liability standards do not. While our measures of resource-based regulatory intensity do not predict every financial outcome equally well, they predict enough of them well to support the view that public regulation plays an important role in setting the institutional foundations for securities markets.

D. Corruption and Enforcement

Skepticism about public enforcement, particularly in developing nations, may be due to the well-founded view that government officials — the market regulators themselves — often operate as a “grabbing hand” rather than a “helping hand” in the matters of economic regulation. Shleifer and Vishny (1999). In developing nations, neither corrupt regulators who use their public power to extract bribes nor regulators in overly statist governments that use regulation to punish political opponents will build

financial markets. Building up public enforcement there may just further empower deleterious elements in that society.¹⁴ Hence, analysts convinced of such a view of governmental regulatory misdirection would prefer private litigation, which they see as out from the control of government's grabbing hand.

But there's little reason to believe that *private* litigation via the judiciary in such nations is structurally *more* efficacious than *public* regulation: one would expect *both* the judiciary and the regulators to be corrupt and ineffective. For both public and private enforcement, the problem confronting nations with corrupt governments is their need to improve public administration, improving either regulatory bodies in the case of public enforcement or the courts in the case of private enforcement. It is not obvious which institutional structure — the judiciary or the regulators — is easier to reform. Indeed, since courts tend to deal with a broad array of issues (contract, tort, family law, and so on), those seeking to strengthen financial markets in such nations may find it *easier* to build up specialized regulators first. The mechanisms for providing such specialized assistance — through technical assistance from agencies such as the SEC or through financial industry sponsored initiatives — are probably better developed and more effective than more broadly based assistance programs for overall judicial reforms.

E. Developing Better Measures of Public Enforcement

Finally, we note the importance of developing better measures of public enforcement. Budgets and staffing levels for a number of jurisdictions were not available and it would be preferable to gather direct information about the allocation of staffing and budgetary resources in countries with consolidated regulatory operations rather than relying on the extrapolating methods that we use. It would, moreover, be useful to collect information on the actual enforcement activities undertaken in each jurisdiction: how many cases prosecuted per year; how many sanctions imposed and with what level of monetary penalty; how many criminal convictions obtained, as the World Bank has successfully accomplished in related areas. It would also be useful to collect information on regulatory resources dedicated to other public

¹⁴ The La Porta et al. (2006) conclusions about public enforcement seem to have been influenced by the fact that in less developed countries their formal public enforcement variable was less strongly associated with robust capital markets than in their sample as a whole. This distinction did not exist for our public enforcement variables; indeed, the relationship between our public enforcement variables and capital market size is even stronger when we limit the analysis to non-OECD countries, although this sub-sample is small for conventional statistical analysis.

functions, such as promulgating new regulations or conducting onsite examinations. If this paper does nothing more than to stimulate the World Bank and other agencies to develop more authoritative data, we will see our efforts as successful. Until then, the data does not support a conclusion that public enforcement is less valuable than private enforcement in developing financial markets.

Conclusion

The relative value of public and private enforcement is a key issue for understanding how financial markets deepen and broaden. Recent work in finance has focused on how securities laws work, concluding that private enforcement via private lawsuits among contracting parties works best. Public enforcement is seen as working poorly. Yet the legal literature has persistently found deep flaws when analyzing private enforcement mechanisms prevailing in the United States. Structural elements — like dispersed ownership — make collective private remedies potentially quite suboptimal and put authority in the lawyers' hands, an authority that we know is not relentlessly used in outside investors' interests.

Given this *a priori* indeterminacy, we constructed two categories of resource-based public enforcement intensity, one using regulatory budgets and another using regulatory staffing. We found significant correlation between measures of public enforcement and key financial outcomes, such as stock market capitalization, trading volumes, the number of domestic firms, and the number of IPOs. Moreover, in horse races between our measure of public enforcement and existing measures of private enforcement, public enforcement typically does no worse than disclosure-based private enforcement (and much better than liability-based private enforcement) in explaining these financial outcomes around the world.

The apparent importance of regulatory intensity seems to go against the basic characteristics of legal origin type. Intensity's importance thus casts some doubt on the literature assessing the importance of legal origin to financial outcomes, since the tool of public enforcement (as opposed to fiduciary-oriented private litigation) has not usually been strongly associated with the common law.

Still, there is more to do. Neither the public enforcement measures nor the private enforcement variables predict several intermediate outcomes associated with the development of strong capital markets, such as stock market efficiency or other more technical measures of stock market performance.

We emphasize that we do not see the data as telling us that public enforcement is more important than private enforcement. First, some financial indicators — in particular, those associated with dispersed ownership — tie more strongly to private enforcement indices than to our public enforcement measures. Second, even where public enforcement is statistically significant, some disclosure-based private indicators are often significant as well. Third, the measures of both public and private enforcement — and of their effectiveness — are highly imperfect. Further improving how we measure these is possible for both and may yield better understanding of which outcomes public enforcement most affects, which ones private enforcement influences, and how the two main enforcement mechanisms interact.

Overall, and most importantly, we caution against using the current explanations for the relative power of private and public enforcement in making public policy around the world. Public enforcement as we measure it does rather well in the regressions. Since public actors typically design both the private and public systems, the debilities of public actors can creep into both designs. The evidence does not warrant rejecting public enforcement in favor of private enforcement for building good securities markets.

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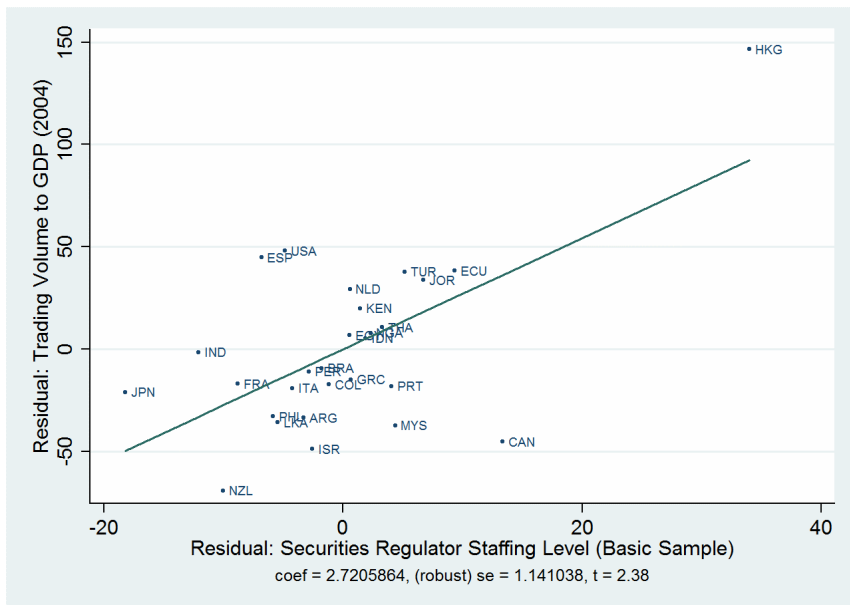
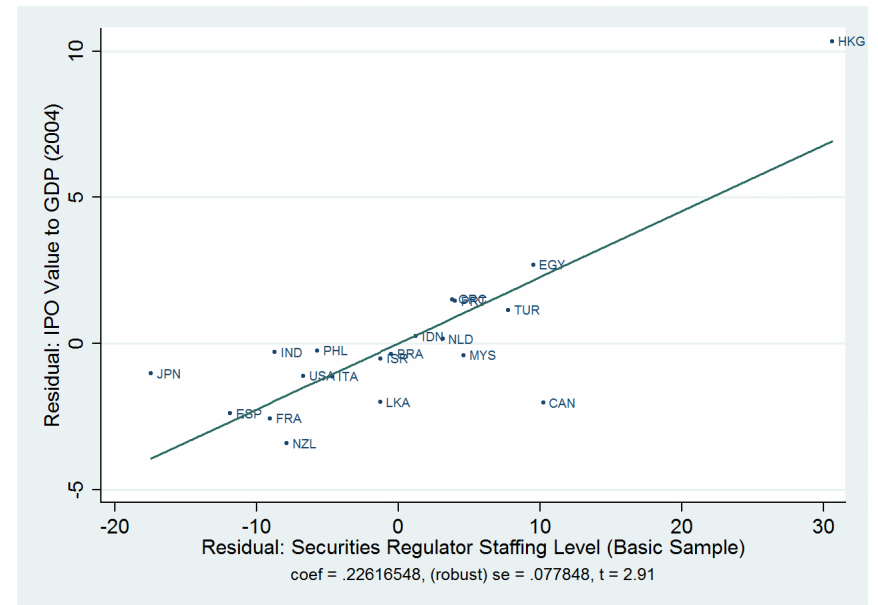
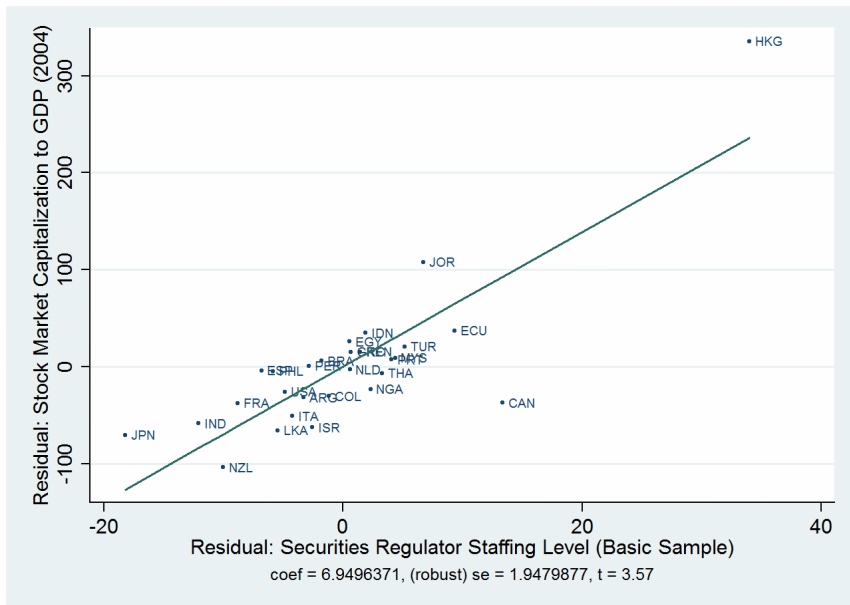


Figure 1. Residual plot of financial market size on securities regulator staffing (basic sample)

These figures plot the residuals of four financial market outcomes on the residual for the regulators' staffing levels (basic sample). The model specification is of the form in Table 5, column 9, but using the basic budget sample. Independent variables are disclosure, liability standards, anti-director rights, a wealth control and judicial efficiency.

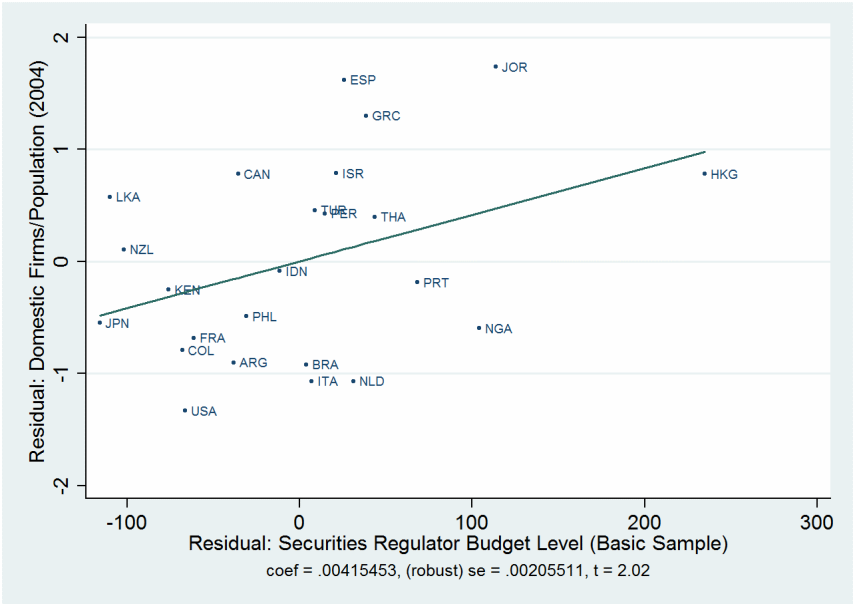
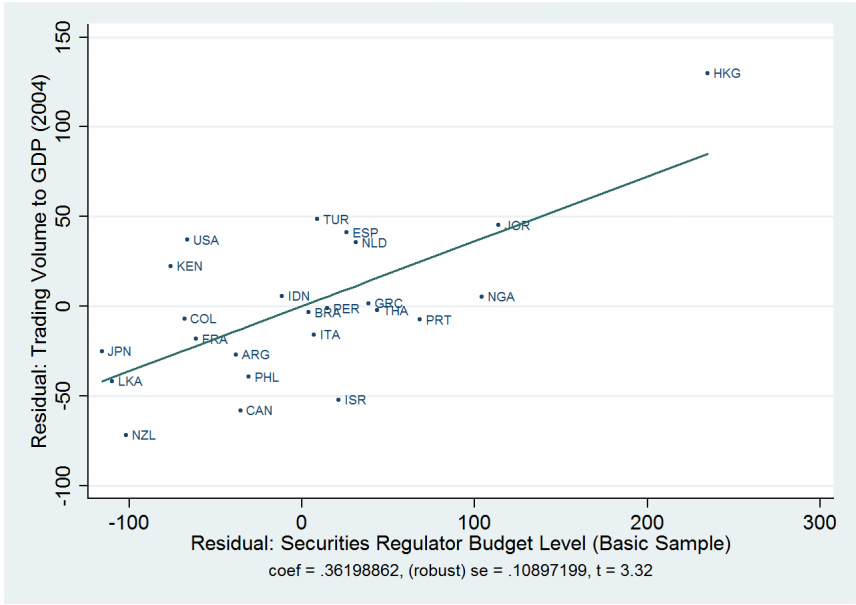
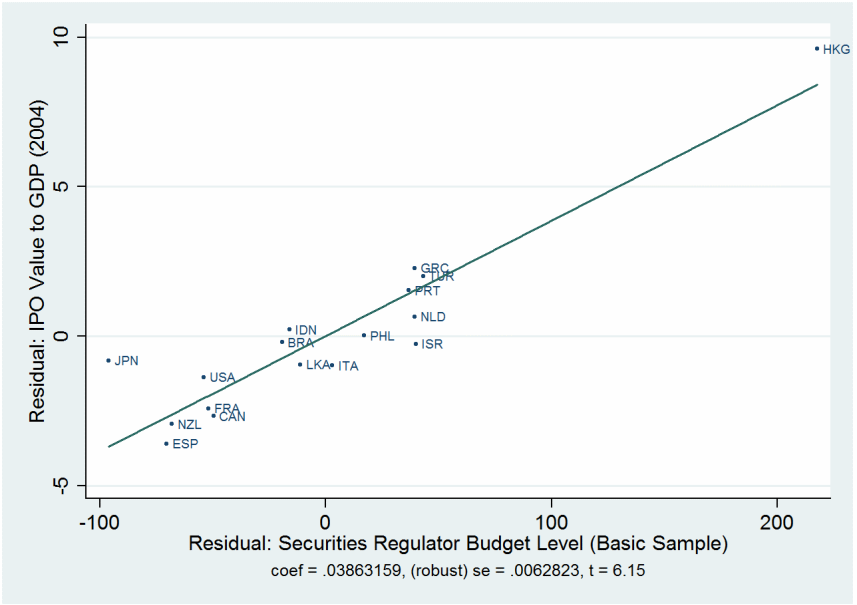
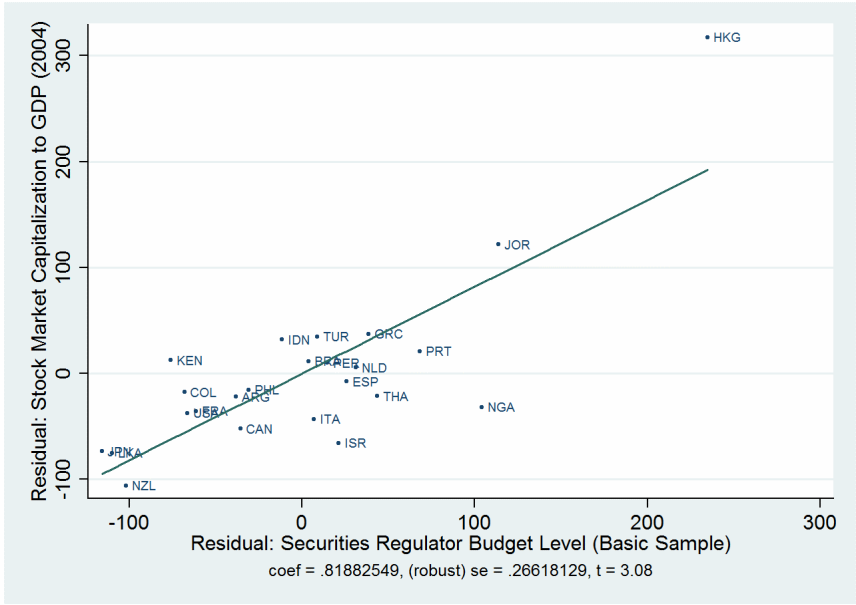


Figure 2. Residual plot of financial market size on securities regulator budget (basic sample)

These figures plot the residuals of four financial market outcomes on the residual for the regulators' budget levels (basic sample). The model specification is of the form in Table 6, column 5, but using the basic budget sample. Independent variables are disclosure, liability standards, anti-director rights, a wealth control and judicial efficiency.

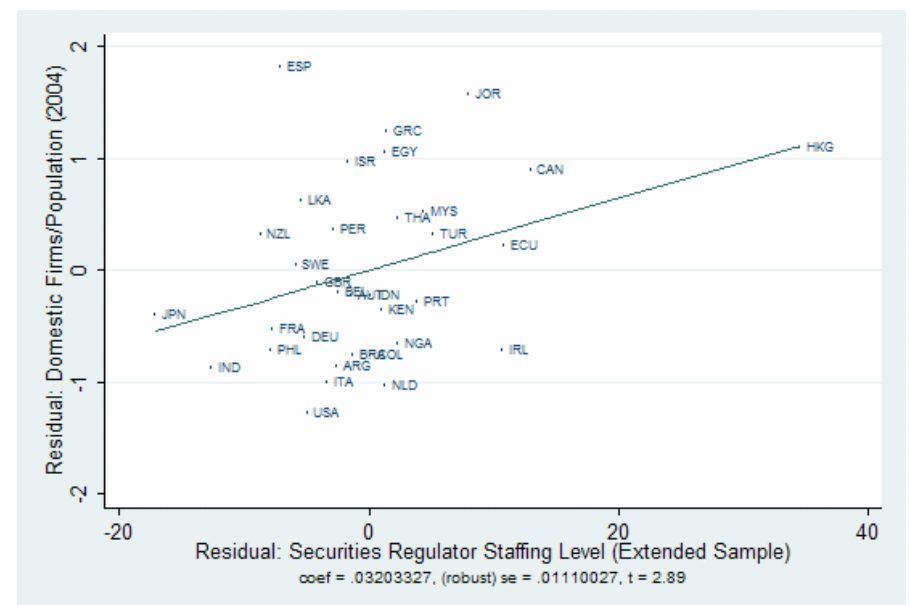
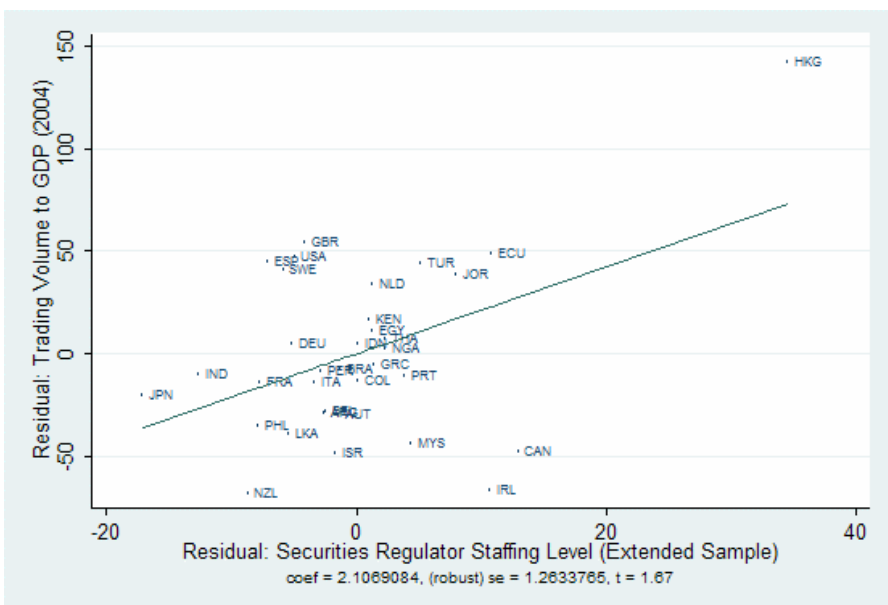
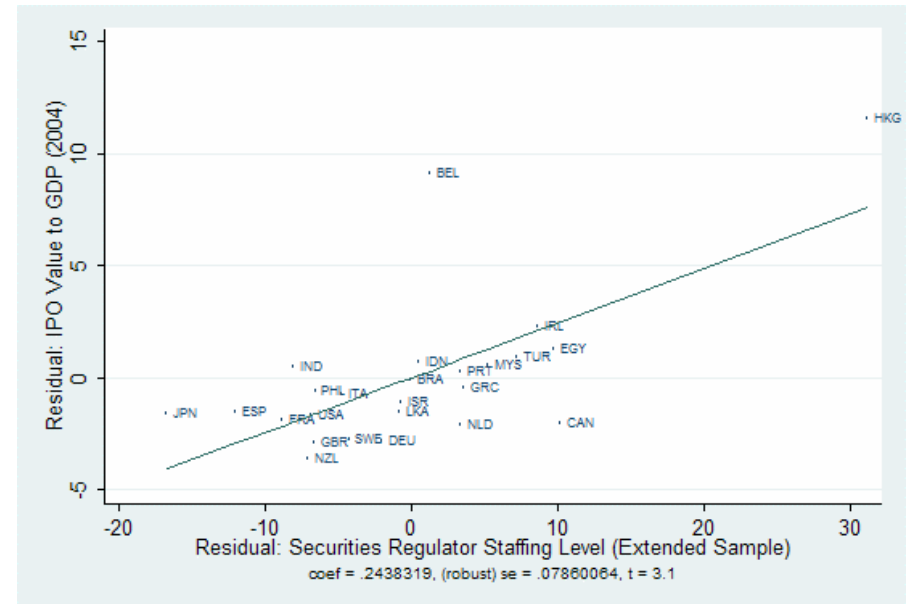
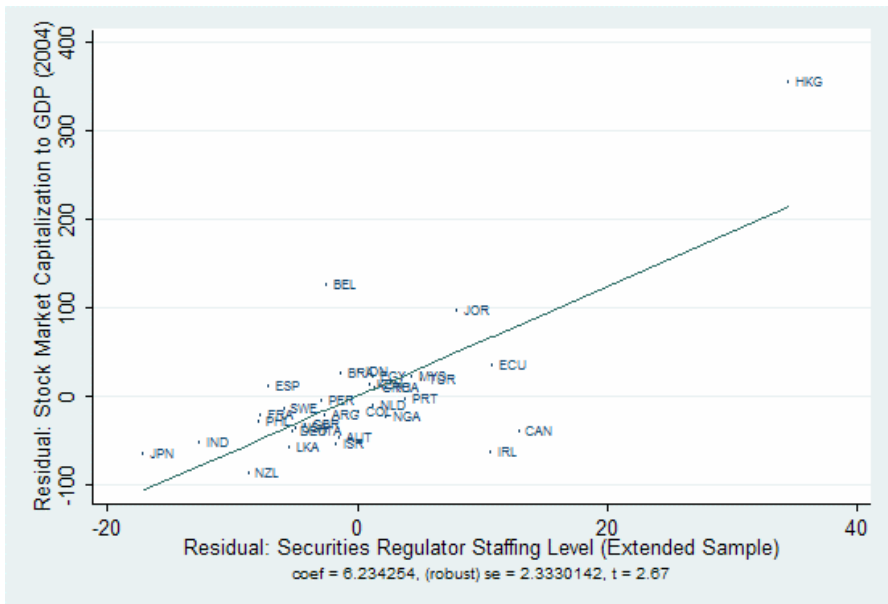


Figure 3. Residual plot of financial market size on securities regulator staffing (extended sample)

These figures plot the residuals of four financial market outcomes on the residual for the regulators' staffing levels (extended sample). The model specification is of the form in Table 5, column 9. Independent variables are disclosure, liability standards, anti-director rights, a wealth control and judicial efficiency.

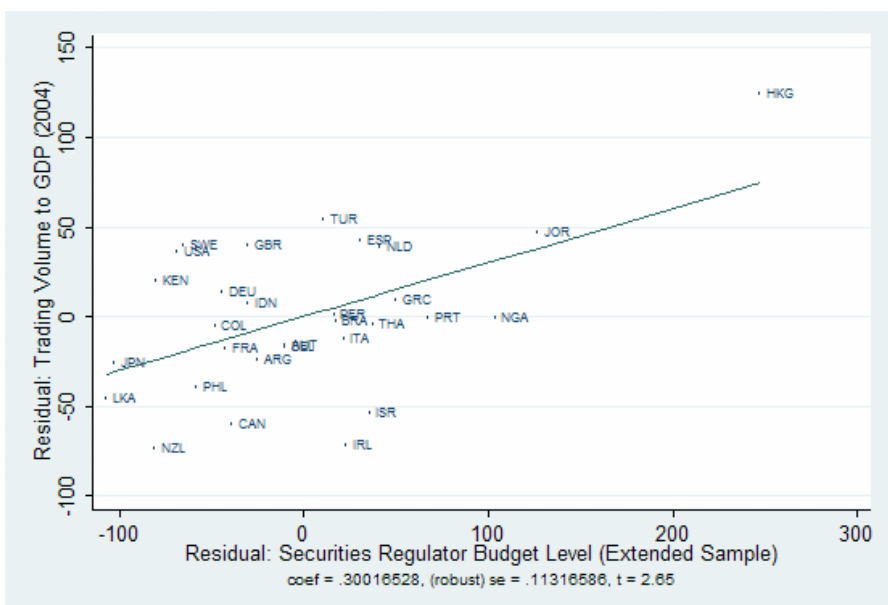
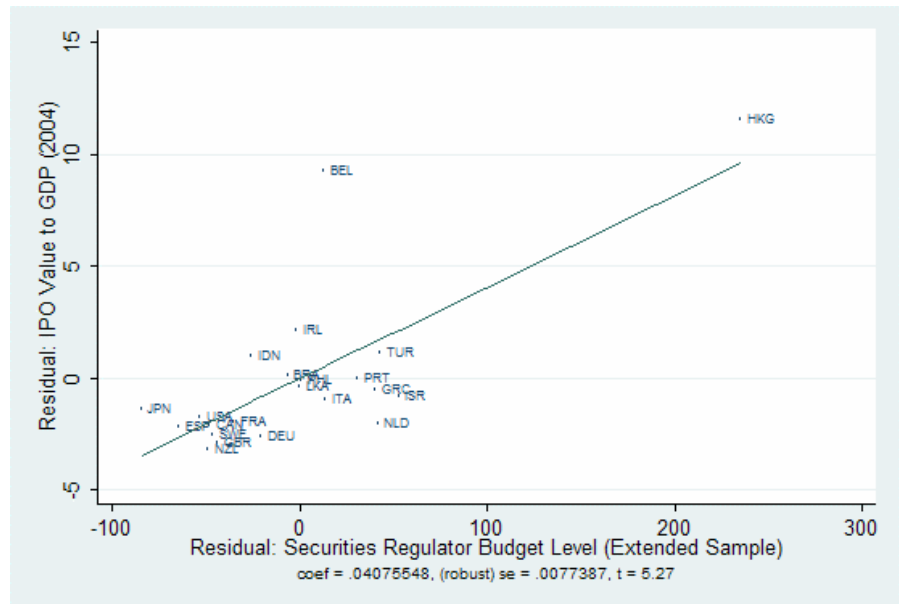
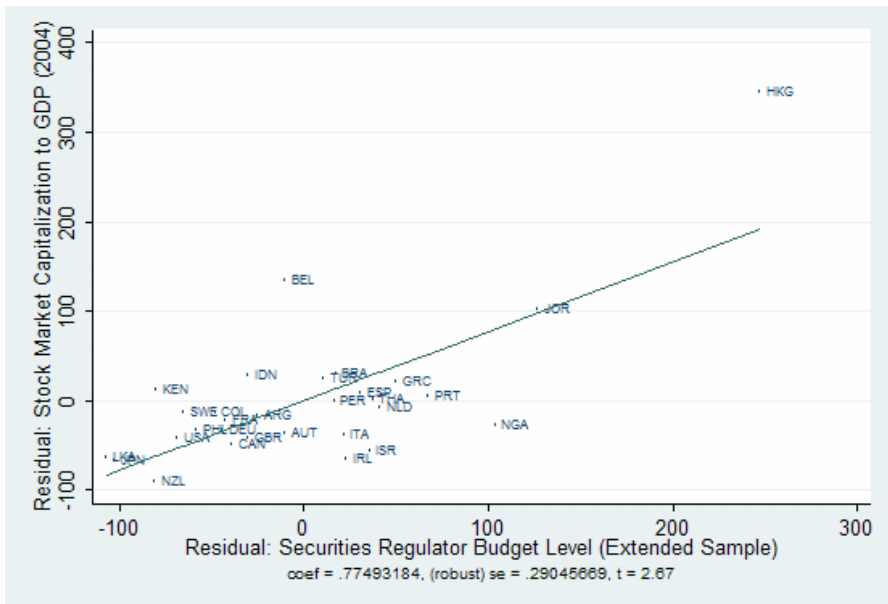


Figure 4. Residual plot of financial market size on securities regulator budget (extended sample)

These figures plot the residuals of four financial market outcomes on the residual for the regulators' budget levels (extended sample). The model specification is of the form in Table 6, column 5. Independent variables are disclosure, liability standards, anti-director rights, a wealth control and judicial efficiency.

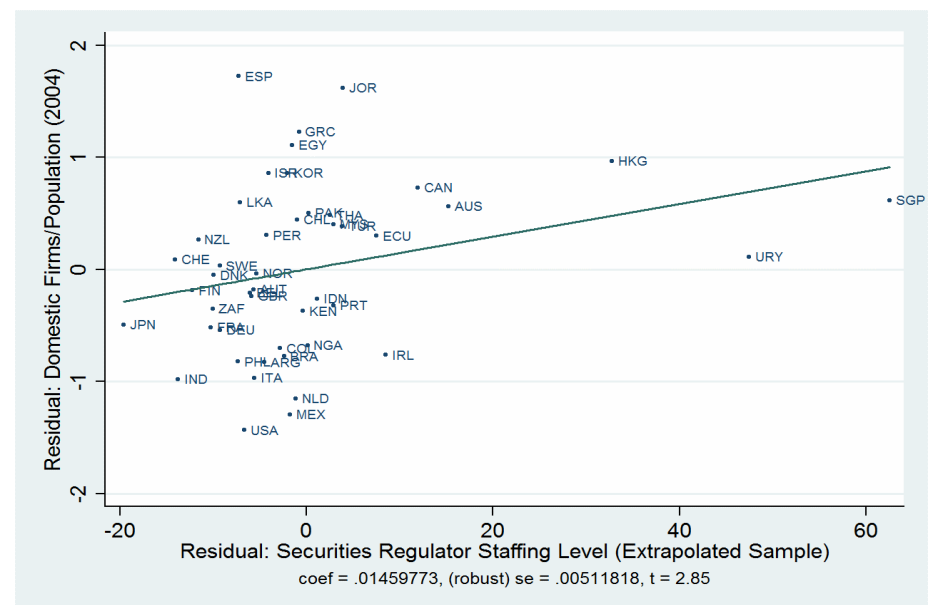
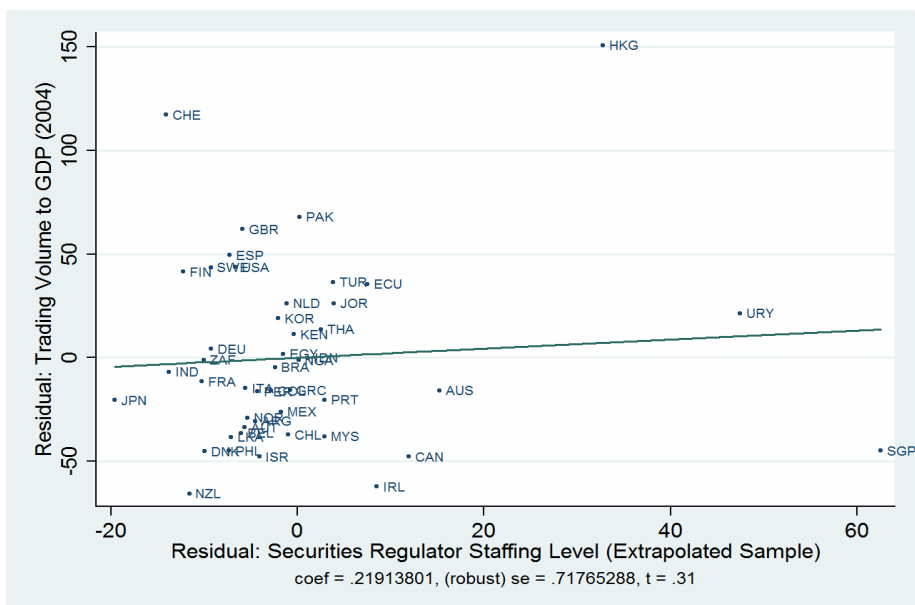
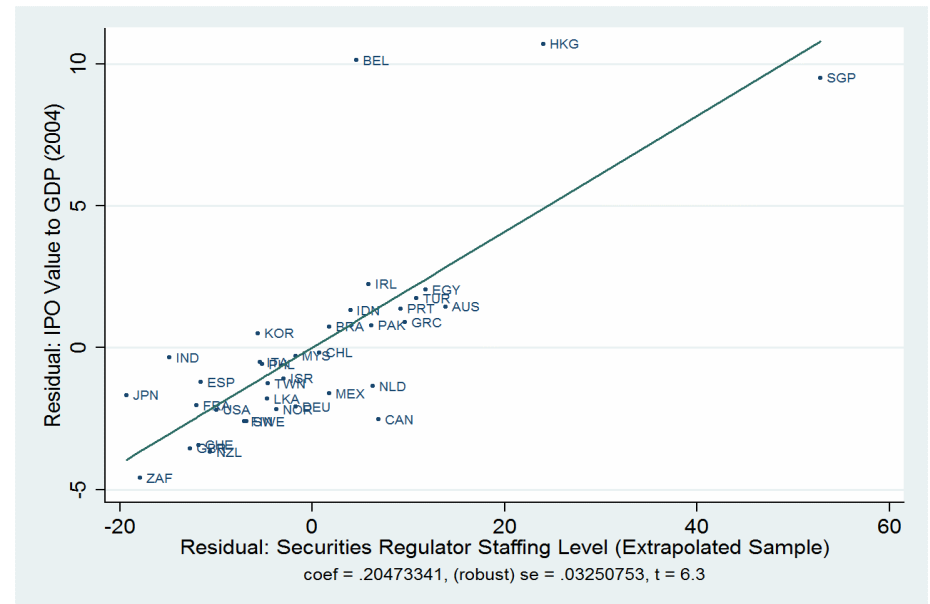
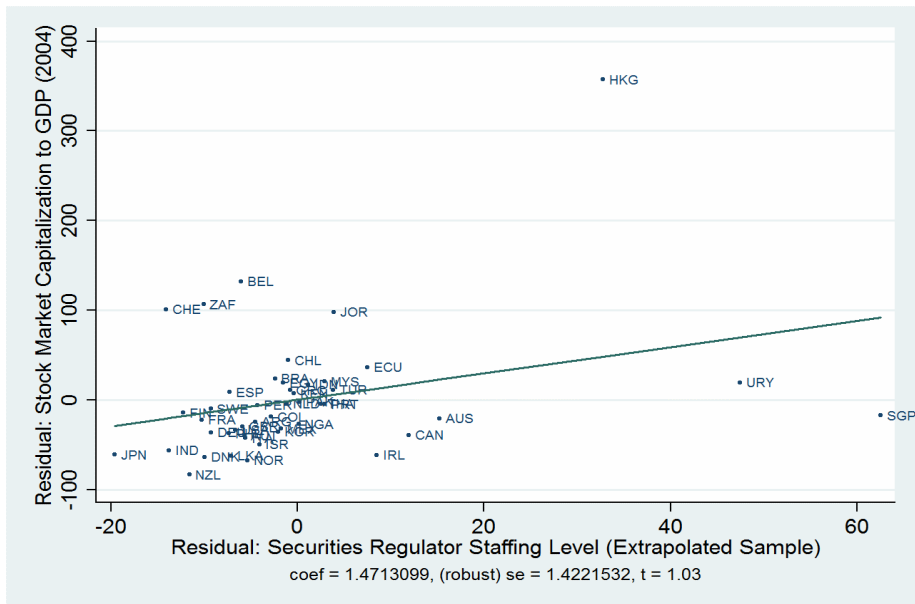


Figure 5. Residual plot of financial market size on securities regulator staffing (extrapolated sample)

These figures plot the residuals of four financial market outcomes on the residual for the regulators' staffing levels (extrapolated sample). The model specification is of the form in Table 5, column 9, but using the extrapolated staffing sample. Independent variables are disclosure, liability standards, anti-director rights, a wealth control and judicial efficiency.

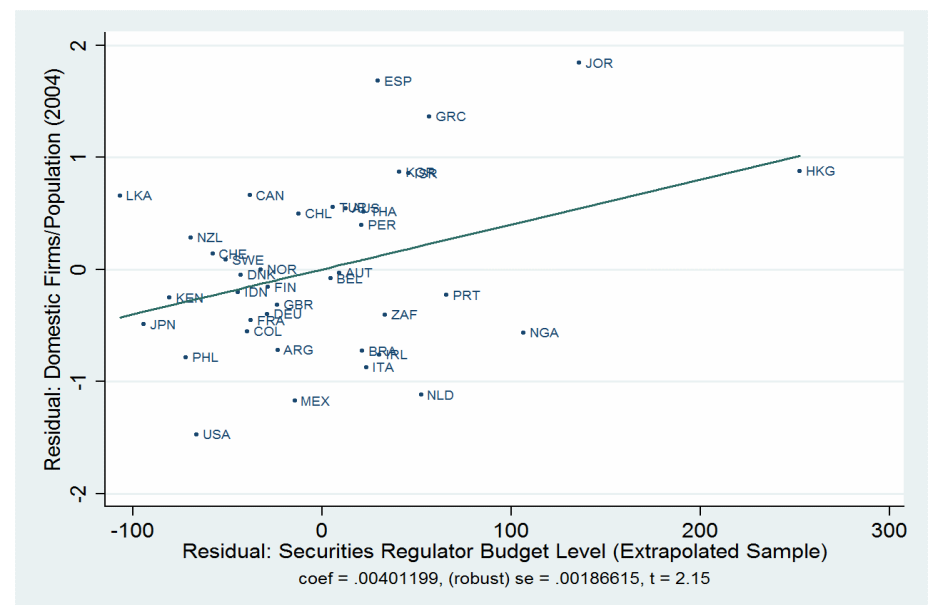
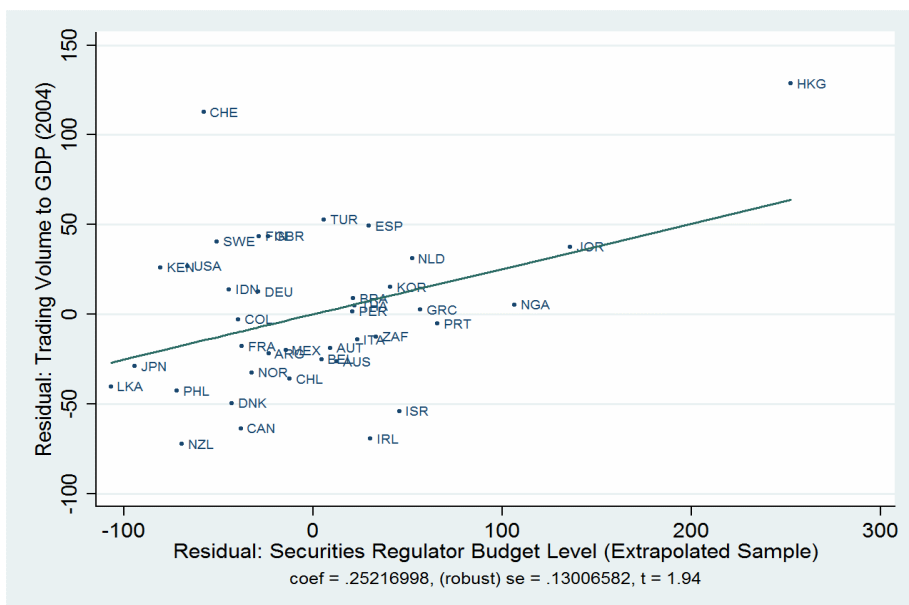
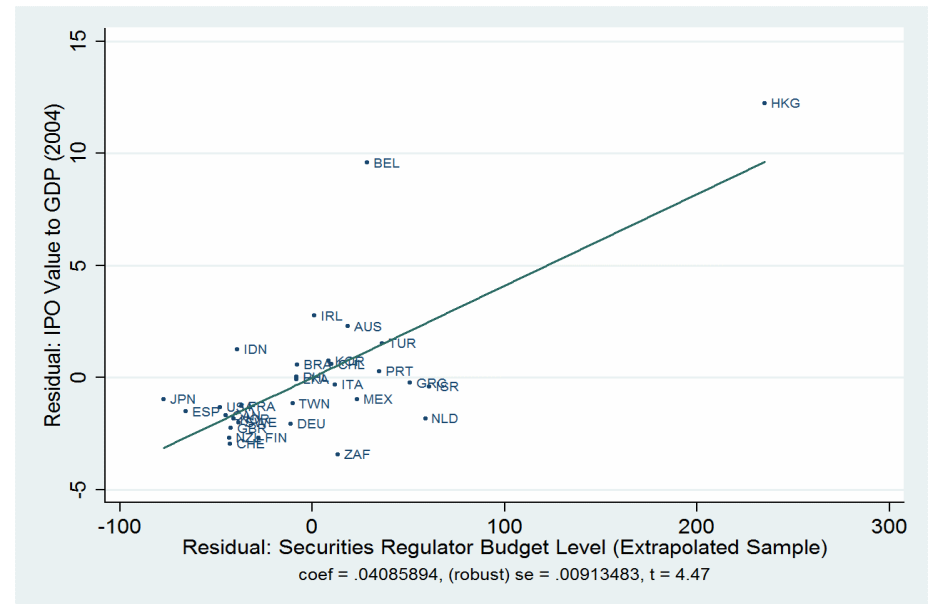
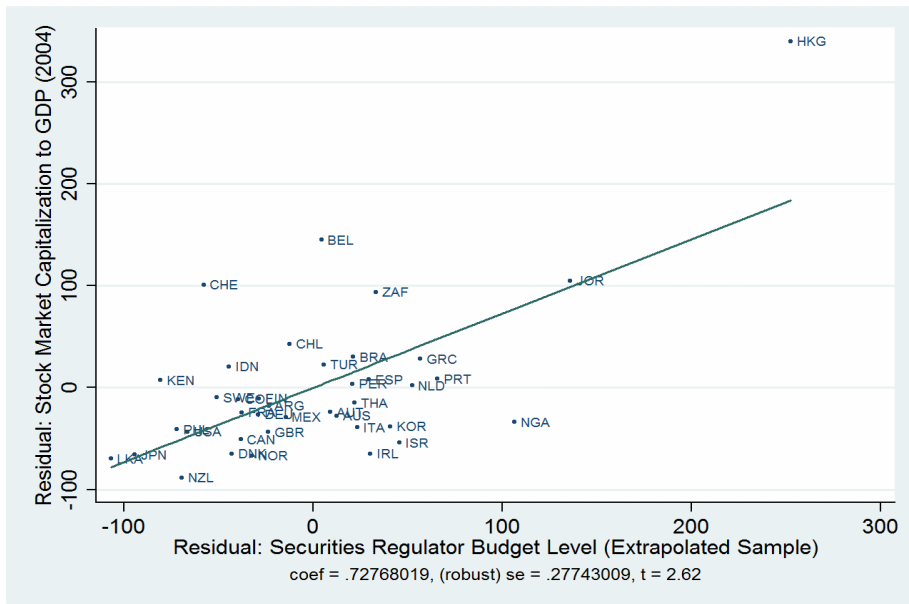


Figure 6. Residual plot of financial market size on securities regulator budget (extrapolated sample)

These figures plot the residuals of four financial market outcomes on the residual for the regulators' budget levels (extrapolated sample). The model specification is of the form in Table 6, column 5, but using the extrapolated budget sample. Independent variables are disclosure, liability standards, anti-director rights, a wealth control and judicial efficiency.

Table 1. Description of variables.

This table lists our independent, dependent, and control variables. We measure public enforcement first with resources and second via the regulator's formal powers. Resources are measured by staffing levels (per million of population) and budgets (per billion US of GDP). For each resource measure, we begin with a basic sample, as reported by the How Countries Supervise (2006) databook. We extend each sample with staffing and budget data from other sources. We also extrapolate securities regulators' budgets and staffing for integrated regulators. We use public enforcement measures that look to the enforcement authorities' formal powers. Private enforcement variables measure the intensity of disclosure rules and measure the nature of liability in private litigation. The control variables seek to measure the countries' legal environment, their per capita wealth and income, and in some regressions, the nations' legal origin. Outcomes are standard measures of the size of securities markets, ownership considerations, and market efficiency.

Variable	Description	Sources
<i><u>Public enforcement: resource-based and formal powers</u></i>		
Staff per million population (basic sample)	The 2005 size of the securities regulators' staff, divided by the country's population in millions. The data is limited to those countries who report their staffing levels in the semi-official Guide to Financial Regulators, which reports 30 countries' staffing levels.	How Countries Supervise (2006); including only those countries where actual staffing levels were reported. Population data downloaded from World Bank Data and Statistics website.
Staff per million population (extended sample)	The 2005 size of the securities regulators' staff, divided by the country's population in millions. This data adds to the basic sample of Staff per million population observations for countries that report staffing data for a consolidated financial services regulator and for which the regulators' annual reports or similar official documents report the relevant staffing data. This adds observations for Belgium, Germany, Iceland, Ireland, Luxembourg, and the United Kingdom, yielding 36 observations in total.	How Countries Supervise (2006); regulators' annual reports and similar for several nations. Population data downloaded from World Bank Data and Statistics website. Individual country online sources are listed in the appendix.
Staff per million population (extended and extrapolated sample)	The 2005 size of the securities regulators' staff, divided by the country's population in millions. For this variable we added to Extended Staff observations for countries with consolidated financial supervisory bodies that do not report separate staffing levels for securities regulators. We extrapolated securities staffing estimates by using the median ratios of securities regulator staffing to the consolidated regulator's staffing for the jurisdictions that reported enough data to construct such a ratio. This yields 53 observations.	How Countries Supervise (2006); regulators' annual reports and similar official documents for the additional nations. Population data downloaded from World Bank Data and Statistics website.
Budget per billion US\$ of GDP (basic sample)	As with the basic Staff per million population, this variable is the securities' regulators' 2005 budget divided by the country's GDP, limited to those countries that report their staffing levels in the semi-official Guide to Financial Regulators, which reports 26 observations here. The budget reports were adjusted to US dollars at the December 31, 2005 exchange rates.	How Countries Supervise (2006); GDP data downloaded from World Bank Data and Statistics website.
Budget per billion US\$ of GDP (extended sample)	Analogous to extended Staff sample, yielding 32 observations, by adding budgets for the securities regulator in a consolidated financial authority for which objective allocation of the securities regulatory budget was possible. Budget in local currency were adjusted to US dollars at the December 31, 2005 exchange rate.	How Countries Supervise (2006); regulators' annual reports and similar official documents. GDP data downloaded from World Bank Data and Statistics website.
Budget per billion US\$ of GDP (extended and extrapolated sample)	Analogous to extrapolated Staff sample, yielding 46 observations. We added 14 observation to the extended budget sample for nations where overall staffing of the consolidated financial agency was available. We used the median ratio of securities staff to consolidated staff for countries in which the actual staffing levels of each were available, and the overall staffing level for each of these 14 observations by that ratio. Budgets in local currency were adjusted to US dollars at the December 31, 2005 exchange rate.	How Countries Supervise (2006); regulators' annual reports and similar official documents. GDP data downloaded from World Bank Data and Statistics website.
Formal public enforcement index (LLS)	Formal public enforcement equals the arithmetic mean of: (1) supervisor characteristics index; (2) its rule-making power index; (3) its investigative powers index; (4) orders authority index; and (5) criminal authority index, as described in La Porta et al. (2006).	La Porta et al. (2006)
Formal public enforcement index (Djankov)	Public enforcement here is an index aggregating whether certain suspect corporate transactions can lead to a fine or jail sentences for the approving body, or fine or jail sentence for the principal wrongdoer.	Djankov et al. (2005)
<i><u>Private enforcement indices</u></i>		
Disclosure index	The index of disclosure equals the arithmetic mean of: (1) nature of liability on a prospectus; (2) extent compensation must be disclosed; (3) shareholders disclosure; (4) extent inside ownership must be disclosed; (5) extent irregular contracts must be disclosed; (6) and the extent that related party and irregular transactions must be disclosed, as described in La Porta et al. (2006).	La Porta et al. (2006)
Liability standards index	The index of liability standards equals the arithmetic mean of: (1) liability standard for the issuer and its directors; (2) liability standard for the distributor; and (3) liability standard for the accountant, as described in La Porta et al. (2006).	La Porta et al. (2006)

Table 1 (cont'd). Description of variables.

<u>Control variables</u>		
Anti-director rights index (original and revised)	An index formed by adding one when: (1) shareholders can mail in their proxy vote; (2) shareholders are not required to deposit their shares prior to the shareholders' meeting; (3) cumulative voting is allowed; (4) an oppressed minorities mechanism is in place; (5) shareholders have preemptive rights by default; and (6) shareholders owning or voting 10% of the share capital can call a shareholders' meeting.	La Porta et al. (1998) (original), revised in Djankov et al. (2005)
Judicial efficiency index	Assessment of the efficiency and integrity of the legal environment as it affects business, particularly foreign firms produced by the country risk rating agency International Country Risk (ICR). It may be taken to represent investors assessment of conditions in the country in question. Average between 1980 and 1983. Scale from 0 to 10, with lower scores representing lower efficiency levels.	Political Risk Services (1996); International Country Risk Guide
Common law jurisdiction	Coding of countries by legal origin, as provided by LLSV (1996). Any data missing from La Porta et al. is coded using the CIA Factbook.	LLSV (1996: 1130-31), CIA Factbook (2006)
Ln GNI per capita	The log of gross national income per capita in 2005 adjusted on the basis of purchasing power parity.	Derived from data downloaded from World Bank Data and Statistics website.
Log of GDP	The log of 2005 gross domestic product in current US dollars.	Derived from GDP variables downloaded from World Bank Data and Statistics website.
<u>Outcome variables</u>		
Market capitalization	Ratio of a country's 2004 stock market capitalization to its GDP	World Bank (2006b)
Trading volume	Ratio of a country's 2004 stock market trading volumes to GDP	World Bank (2006b)
Domestic firms	The ratio of the number of a country's listed domestic firms in 2004 to millions of population	World Bank (2006b)
IPOs to GDP	The ratio of annual capital raised in a country through IPOs in 2004 to its GDP	World Bank (2006b)
Block premia	The block premia is computed by taking the difference between the price per share paid for the control block and the exchange price two days after the announcement of the control transaction, dividing by the exchange price and multiplying by the ratio of the proportion of cash flow rights represented in the controlling block, as described in Dyck and Zingales (2004).	Dyck and Zingales (2004)
Ownership concentration	Average percentage of common shares not owned by the top three shareholders in the ten largest non-financial, privately-owned domestic firms in a given country.	La Porta et al. (2006).
Access to equity	Index of the extent to which business executives in a country agree with the statement "Stock markets are open to new firms and medium-sized firms". Scale from 1 (strongly agree) though 7 (strongly disagree)	Schwab et al. (1999)
Equity index	Index averaging other indices that measure market size, market access, market efficiency, and market stability in 2004.	World Bank (2006b)
Equity Market Size Index	Index of factors measuring the size of capital markets in 2004	World Bank (2006b)
Equity Market Access Index	Index of factors measuring the ease of accessing capital markets in 2004	World Bank (2006b)
Equity Market Efficiency Index	Index of factors measuring the efficiency of capital markets in 2004	World Bank (2006b)
Equity Market Stability Index	Index of factors measuring the stability of capital markets in 2004.	World Bank (2006b)

Table 2. Resource-based securities law enforcement data: staffing and budget.

This table provides resource-based data to measure public enforcement intensity based on securities regulators' staffing and budget levels. The basic sample, in column 1, shows the staffing levels per million of a country's population from *How Countries Supervise* (2006). Column 2 extends the sample to countries that provide their securities' regulators budget through other sources (indicated in Table 1). Column 3 adds in extrapolated staffing levels for integrated regulators: Some nations, such as the United Kingdom, have a single financial markets regulator, one that regulates securities markets, banking, and sometimes insurance. For integrated regulators that do not break out their securities regulators' budgets from the integrated budget, we extrapolated the securities regulatory budget using the median proportion of securities regulator budget/integrated regulatory budget for those nations for which we could determine both. Columns 4, 5, and 6 show resource-based securities regulator data for the regulators' budgets, measured analogously. Columns 7 and 8 show previously-assembled measures of public enforcement based on the regulator's formal authority, such as independence from the executive, rule-making authority, authority to bring criminal actions, and so on.

Country	Resource-based measures of public enforcement						Formal measures of public enforcement	
	Staff per million of population			Budget per billion US\$ of GDP			(7) Public Enforcement Index (LLS)	(8) Public Enforcement Index (Djankov)
	(1) Basic observations	(2) Basic sample, with extended observations	(3) Extended sample, with extrapolated observations	(4) Basic observations	(5) Basic sample, with extended observations	(6) Extended sample, with extrapolated observations		
Argentina	3.46	3.46	3.46	\$15,984	\$15,984	\$15,984	0.58	0.00
Australia			39.35			120,162	0.90	0.50
Austria		9.24	9.24		34,497	34,497	0.17	1.00
Belgium		12.93	12.93		67,274	67,274	0.15	0.50
Brazil	2.68	2.68	2.68	35,260	35,260	35,260	0.58	0.50
Canada	38.98	38.98	38.98	83,932	83,932	83,932	0.80	1.00
Chile			10.39			67,137	0.60	1.00
Colombia	3.88	3.88	3.88	46,603	46,603	46,603	0.58	0.00
Czech Republic	12.95	12.95	12.95	41,685	41,685	41,685		1.00
Denmark			12.34			33,745	0.37	0.75
Ecuador	7.41	7.41	7.41				0.55	1.00
Egypt	3.59	3.59	3.59				0.30	0.00
Finland			10.18			48,969	0.32	0.00
France	5.93	5.93	5.93	29,205	29,205	29,205	0.77	0.50
Germany		4.43	4.43		13,527	13,527	0.22	1.00
Greece	12.17	12.17	12.17	79,801	79,801	79,801	0.32	0.50
Hong Kong	58.47	58.47	58.47	365,740	365,740	365,740	0.87	0.00
Hungary		10.90	10.90		82,120	82,120		0.00
Iceland		15.86	15.86		44,094	44,094		0.00
India	0.43	0.43	0.43				0.67	0.50
Indonesia	1.97	1.97	1.97	5,571	5,571	5,571	0.62	0.00
Ireland		28.02	28.02		88,871	88,871		0.00
Israel	18.82	18.82	18.82	153,118	153,118	153,118	0.63	1.00
Italy	7.40	7.40	7.40	62,889	62,889	62,889	0.48	0.00
Japan	4.61	4.61	4.61	17,000	17,000	17,000	0.00	0.00
Jordan	14.90	14.90	14.90	326,569	326,569	326,569	0.60	0.00
Kenya	0.96	0.96	0.96	93,953	93,953	93,953	0.70	0.00
Korea			13.15			95,147	0.25	0.50
Luxembourg		314.43	314.43		512,897	512,897		1.00
Malaysia	22.65	22.65	22.65				0.77	1.00
Mexico			4.71			52,494	0.35	0.50
Netherlands	23.52	23.52	23.52	138,785	138,785	138,785	0.47	0.00
New Zealand	9.00	9.00	9.00	37,784	37,784	37,784	0.33	0.00
Nigeria	4.60	4.60	4.60	338,349	338,349	338,349	0.33	0.00
Norway			15.99			34,106	0.32	1.00
Pakistan			2.47				0.58	0.75
Peru	5.11	5.11	5.11	109,651	109,651	109,651	0.78	0.25
Philippines	4.37	4.37	4.37	66,120	66,120	66,120	0.83	0.00
Poland	4.64	4.64	4.64	22,961	22,961	22,961		1.00
Portugal	14.49	14.49	14.49	80,842	80,842	80,842	0.58	1.00
Singapore			88.28				0.87	1.00
Slovak Republic			13.93			53,844		0.00
South Africa			3.82			118,453	0.25	0.00
Spain	7.33	7.33	7.33	29,931	29,931	29,931	0.33	0.75
Sri Lanka	2.35	2.35	2.35	46,126	46,126	46,126	0.43	0.00
Sweden		10.23	10.23		27,234	27,234	0.50	1.00
Switzerland			8.04			31,418	0.33	0.50
Taiwan			14.65			73,061	0.52	0.00
Thailand	6.40	6.40	6.40	83,804	83,804	83,804	0.72	0.00
Turkey	6.18	6.18	6.18	45,417	45,417	45,417	0.63	0.00
United Kingdom		19.05	19.05		81,709	81,709	0.68	0.00
United States	23.64	23.29	23.29	76,459	76,459	76,459	0.90	0.00
Uruguay			50.66				0.57	0.50
Venezuela							0.55	0.00
Zimbabwe							0.42	
Number of observations	30	39	53	26	35	46	49	54
Mean	11.09	19.43	19.73	\$93,598	\$96,736	\$89,441	0.52	0.40
Median	6.29	7.41	10.18	\$64,504	\$66,120	\$64,504	0.55	0.38
Minimum	0.43	0.43	0.43	\$5,571	\$5,571	\$5,571	0.00	0.00
Maximum	58.47	314.43	314.43	\$365,740	\$512,897	\$512,897	0.90	1.00
Standard deviation	12.48	49.79	44.19	\$99,084	\$113,586	\$100,764	0.22	0.43

Table 3. Pairwise Correlation Matrix for Key Enforcement Variables.

In this table we report pairwise correlations between the key enforcement variables. The six resource-based measures of enforcement correlate at levels of .54 and higher. The resource-based measures of public enforcement do not correlate as strongly with enforcement measures derived from the regulators' formal powers, with the correlations at levels ranging from -.05 to .41. Of note is that the two public enforcement measures based on formal powers of the regulator correlate negatively.

	<i>Disclosure</i>	<i>Liability</i>	<i>Anti-director rights (original)</i>	<i>Judicial efficiency</i>	<i>Public enforcement (LLS)</i>	<i>Public enforcement (Djankov)</i>	<i>Regulatory staff (basic)</i>	<i>Regulatory staff (extended)</i>	<i>Regulatory staff (extrapolated)</i>	<i>Regulatory budget (basic)</i>	<i>Regulatory budget (extended)</i>	<i>Regulatory budget (extrapolated)</i>
Disclosure	1.00											
Liability standards	0.55	1.00										
Anti-director rights	0.52	0.50	1.00									
Judicial efficiency	0.25	0.22	0.21	1.00								
Public enforcement (LLS)	0.33	0.31	0.37	-0.11	1.00							
Public enforcement (Djankov)	-0.15	-0.08	0.06	0.24	-0.01	1.00						
Regulatory staff (basic)	0.40	0.46	0.40	0.56	0.39	0.17	1.00					
Regulatory staff (extended)	0.40	0.44	0.40	0.51	0.32	0.24	1.00	1.00				
Regulatory staff (extrapolated)	0.26	0.28	0.25	0.42	0.41	0.25	1.00	1.00	1.00			
Regulatory budget (basic)	0.25	0.02	-0.03	0.35	0.15	-0.19	0.54	0.54	0.54	1.00		
Regulatory budget (extended)	0.29	0.12	0.05	0.22	0.23	-0.02	0.54	0.71	0.71	1.00	1.00	
Regulatory budget (extrapolated)	0.29	0.14	0.09	0.14	0.25	-0.05	0.54	0.71	0.71	1.00	1.00	1.00

Table 4. Resource-based enforcement and formal powers-based enforcement as predicting financial market size.

In this table, we test whether measures of the level of a nation's resource-based public enforcement predict the level of a nation's financial market strength. We use two measures here of resource-based public enforcement: the securities regulator's staffing level per million of population and the securities regulator's budget/GDP. Financial outcomes are the country's stock market capitalization/GDP, stock trading volume/GDP, number of domestic firms/population, and number of IPOs/GDP, all standard measures of financial development. We control for judicial efficiency, corporate law (via the anti-director rights index), and GDP/capita. We run five models for each pairing of enforcement and outcomes. In the first model, in column 1 for each of the four pairings, we see whether preexisting formal measures of regulatory powers predict the financial outcomes. In the second and fourth models, in columns 2 and 4, we substitute the resource-based measure of public enforcement. In these four pairings, resource-based enforcement is always positive and significant. In the third and fifth models, in columns 3 and 5, we run both the formal measure of public enforcement and the resource-based measure. Formal public enforcement is significant in these models twice. The resource-based measure is significant in 6 of these mini-horse races. Robust standard errors are in parentheses; the constant is not reported.

	Panel A <i>Dependent variable = Stock market capitalization/GDP</i>					Panel B <i>Dependent variable = Trading volume/GDP</i>				
	Formal public enforcement	Staffing resources	Staffing and formal	Budget resources	Budget and formal	Formal public enforcement	Staffing resources	Staffing and formal	Budget resources	Budget and formal
	(1)	(2)	(3)	(4)	(5)	(1)	(2)	(3)	(4)	(5)
Staffing resources (extended)	–	0.013 ^a (0.00)	0.0089 ^c (0.00)	–	–	–	1.71 ^b (0.66)	1.27 (0.81)	–	–
Budget resources (extended)	–	–	–	0.0014 ^b (0.00)	0.00098 ^b (0.00)	–	–	–	0.22 ^a (0.07)	0.18 ^b (0.08)
Formal public enforcement (LLS)	0.34 ^c (0.20)	–	0.46 ^b (0.22)	–	0.55 ^b (0.20)	39.56 (30.01)	–	45.92 (41.58)	–	53.35 (39.98)
Anti-director rights	0.071 ^b (0.03)	0.067 ^c (0.03)	0.04 (0.03)	0.10 ^a (0.03)	0.057 ^c (0.03)	7.86 (4.73)	10.32 ^c (5.20)	7.46 (5.61)	12.61 ^b (5.27)	8.14 (5.55)
Judicial efficiency	0.052 ^b (0.02)	0.03 (0.02)	0.050 ^b (0.02)	0.01 (0.03)	0.04 (0.03)	-2.10 (5.42)	-1.45 (3.13)	0.48 (4.42)	-6.92 ^c (3.92)	-4.10 (5.31)
Ln GDP per capita	0.10 ^a (0.02)	0.051 ^b (0.02)	0.069 ^a (0.02)	0.14 ^a (0.03)	0.13 ^a (0.03)	21.43 ^a (7.08)	9.07 ^b (4.22)	10.95 ^b (4.44)	25.54 ^a (5.23)	25.82 ^a (5.35)
Observations	49	34	34	30	30	49	34	34	30	30
MSE	0.27	0.22	0.21	0.24	0.22	56.89	34.29	33.91	35.51	34.62
Adjusted R ²	0.48	0.63	0.67	0.57	0.64	0.18	0.48	0.49	0.46	0.49

	Panel C <i>Dependent variable = Domestic firms/population</i>					Panel D <i>Dependent variable = IPOs/GDP</i>				
	Formal public enforcement	Staffing resources	Staffing and formal	Budget resources	Budget and formal	Formal public enforcement	Staffing resources	Staffing and formal	Budget resources	Budget and formal
	(1)	(2)	(3)	(4)	(5)	(1)	(2)	(3)	(4)	(5)
Staffing resources (extended)	–	0.036 ^a (0.01)	0.034 ^a (0.01)	–	–	–	0.13 ^a (0.03)	0.13 ^a (0.04)	–	–
Budget resources (extended)	–	–	–	0.0034 ^c (0.00)	0.0027 (0.0019)	–	–	–	0.010 ^b (0.00)	0.0079 ^c (0.00)
Formal public enforcement (LLS)	0.64 (0.48)	–	0.26 (0.68)	–	0.86 (0.74)	3.72 ^b (1.55)	–	0.19 (2.08)	–	2.18 (2.26)
Anti-director rights	0.18 ^b (0.09)	0.04 (0.09)	0.03 (0.10)	0.15 (0.09)	0.08 (0.12)	0.31 (0.24)	0.35 (0.35)	0.34 (0.34)	0.67 ^c (0.38)	0.49 (0.37)
Judicial efficiency	0.26 ^a (0.08)	0.21 ^a (0.07)	0.22 ^a (0.08)	0.15 ^c (0.09)	0.20 ^b (0.10)	0.04 (0.22)	-0.20 (0.20)	-0.19 (0.25)	-0.35 (0.30)	-0.24 (0.35)
Ln GDP per capita	0.29 ^a (0.11)	0.17 (0.12)	0.18 (0.13)	0.41 ^b (0.16)	0.40 ^b (0.16)	1.22 ^a (0.27)	0.75 ^a (0.25)	0.76 ^a (0.27)	1.54 ^a (0.39)	1.51 ^a (0.39)
Observations	49	34	34	30	30	49	34	34	30	30
MSE	0.66	0.64	0.65	0.70	0.69	2.57	2.29	2.33	2.61	2.63
Adjusted R ²	0.66	0.68	0.67	0.65	0.65	0.34	0.46	0.44	0.34	0.33

a = significant at 1%; b = significant at 5 %; and c = significant at 10 %.

Table 5. Resource-based staffing variable in horse race with private enforcement (panels A and B).

In this table we present horse races between resource-based public enforcement and private enforcement variables. Panel A uses market capitalization/GDP as the outcome variable, Panel B trading volume/GDP as the outcome variable, Panel C the number of domestic firms/population, and Panel D IPOs/GDP. We run 9 models with each of the four outcome variables. Column 1 tests public enforcement alone, then we add disclosure in column 2, liability standards in column 3, and both disclosure and liability standards in column 4. Next we add a resource-based measure of public enforcement, the extended sample of the securities regulators' staffing levels. Columns 5 and 6 run the public enforcement variables, with resource-based enforcement alone first in column 5, and then with measured by the securities regulators' powers, in column 6. Columns 7, 8, and 9 check for the robustness of the resource-based enforcement measure in horse races between public and private enforcement: i.e., the staff-based resources variable runs against disclosure in column 7, against liability in column 8, then against both private enforcement measures in column 9. We add three conditioning variables to each model: wealth, judicial efficiency and anti-directors rights. The resource-based measure is significant in 17 of the 20 models in which it appears; disclosure is significant without resource-based enforcement in the majority of runs, but insignificant in 6 out of the 8 models in which it runs against resource-based-enforcement. Liability standards often turns negative in the face of the resource-based measure of public enforcement, significantly so twice. Robust standard errors are in parentheses; the constant is not reported.

<i>Panel A: Dependent Variable = Market Capitalization to GDP (2004)</i>									
	<u>Without Extended Staffing Variable</u>				<u>With Extended Staffing Variable</u>				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Securities regulator staffing level (extended sample)					6.01 ^b (2.22)	6.05 ^b (2.51)	5.82 ^b (2.34)	6.42 ^b (2.30)	6.23 ^b (2.33)
Formal Public Enforcement	83.17 (69.10)	56.84 (71.62)	76.99 (77.82)	57.30 (77.97)	-4.23 (58.36)				
Disclosure		94.15 ^b (36.86)		94.80 ^b (37.14)			32.22 (35.59)		52.91 (32.50)
Liability Standards			19.54 (34.64)	-2.03 (37.03)				-45.88 (45.57)	-59.20 (44.46)
Anti-Directors Rights (rev.)	16.44 (10.98)	6.48 (11.27)	14.77 (12.16)	6.59 (12.17)	3.29 (8.94)	3.45 (9.26)	0.73 (9.68)	7.03 (10.50)	3.91 (10.62)
Judicial Efficiency	11.21 ^c (5.58)	9.65 ^c (5.17)	11.01 ^c (5.81)	9.66 ^c (5.32)	3.16 (4.39)	2.98 (5.14)	2.78 (4.24)	2.85 (4.47)	2.13 (4.45)
Wealth Control	22.68 ^b (9.83)	20.01 ^b (9.79)	21.68 ^b (10.60)	20.10 ^c (10.40)	-4.31 (9.98)	-4.53 (10.23)	-3.88 (9.72)	-4.19 (10.14)	-3.46 (9.69)
Observations	47	47	47	47	34	34	34	34	34
MSE	75.09	73.92	75.88	74.84	58.50	59.53	59.14	58.41	58.46
Adjusted R Squared	0.26	0.29	0.25	0.27	0.60	0.59	0.59	0.60	0.60

<i>Panel B: Dependent Variable = Trading Volume to GDP (2004)</i>									
	<u>Without Extended Staffing Variable</u>				<u>With Extended Staffing Variable</u>				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Securities regulator staffing level (extended sample)					2.33 ^b (1.13)	1.72 (1.33)	1.99 (1.22)	2.33 ^c (1.24)	2.11 (1.26)
Formal Public Enforcement	51.91 (35.49)	33.03 (33.83)	43.07 (36.58)	29.96 (35.63)	54.99 (43.06)				
Disclosure		67.52 ^a (24.78)		63.17 ^b (25.46)			59.45 ^b (27.86)		65.32 ^b (25.62)
Liability Standards			27.93 (25.94)	13.56 (26.72)				-0.36 (34.63)	-16.79 (31.29)
Anti-Directors Rights (rev.)	13.21 ^b (5.12)	6.07 (6.07)	10.82 ^c (6.17)	5.37 (6.60)	14.25 ^b (6.79)	12.22 (7.20)	9.53 (7.08)	14.28 ^c (7.92)	10.43 (7.83)
Judicial Efficiency	4.55 (3.54)	3.43 (3.22)	4.26 (3.67)	3.36 (3.31)	0.70 (3.78)	2.96 (4.67)	-0.01 (3.37)	0.69 (3.90)	-0.19 (3.35)
Wealth Control	25.46 ^a (6.43)	23.55 ^a (6.09)	24.02 ^a (6.58)	22.97 ^a (6.22)	16.21 ^b (6.76)	19.04 ^b (7.02)	17.00 ^b (6.65)	16.21 ^b (6.87)	17.12 ^b (6.75)
Observations	47	47	47	47	34	34	34	34	34
MSE	47.33	46.20	47.50	46.68	42.46	41.87	41.31	43.21	41.87
Adjusted R Squared	0.36	0.39	0.35	0.37	0.50	0.52	0.53	0.49	0.52

a = significant at 1%; b = significant at 5 %; and c = significant at 10 %.

Table 5. Resource-based staffing variable in horse race with private enforcement (panels C and D).

Panel C: Number of domestic firms/population (2004)

	<u>Without Extended Staffing Variable</u>				<u>With Extended Staffing Variable</u>				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Securities regulator staffing level (extended sample)					0.04 ^a (0.01)	0.04 ^a (0.01)	0.03 ^a (0.01)	0.04 ^a (0.01)	0.03 ^a (0.01)
Formal Public Enforcement	0.61 (0.53)	0.26 (0.53)	0.37 (0.56)	0.14 (0.54)		-0.41 (0.80)			
Disclosure		1.25 ^c (0.74)		1.08 (0.80)			0.88 (0.70)		0.88 (0.78)
Liability Standards			0.77 (0.61)	0.53 (0.64)				0.23 (0.56)	0.01 (0.63)
Anti-Directors Rights (rev.)	0.26 ^c (0.13)	0.12 (0.18)	0.19 (0.15)	0.10 (0.18)	0.12 (0.20)	0.14 (0.21)	0.05 (0.22)	0.10 (0.21)	0.05 (0.22)
Judicial Efficiency	0.26 ^a (0.09)	0.24 ^a (0.08)	0.25 ^a (0.09)	0.24 ^a (0.08)	0.19 ^b (0.08)	0.17 (0.10)	0.18 ^b (0.08)	0.19 ^b (0.09)	0.18 ^b (0.08)
Wealth Control	0.56 ^a (0.17)	0.51 ^a (0.18)	0.52 ^a (0.17)	0.49 ^a (0.17)	0.39 ^c (0.22)	0.37 (0.22)	0.40 ^c (0.22)	0.39 ^c (0.23)	0.40 ^c (0.22)
Observations	48	48	48	48	34	34	34	34	34
MSE	0.81	0.79	0.80	0.79	0.83	0.84	0.82	0.84	0.84
Adjusted R Squared	0.62	0.64	0.63	0.64	0.60	0.59	0.61	0.59	0.59

Panel D: Dependent Variable = IPO value/GDP (2004)

	<u>Without Extended Staffing Variable</u>				<u>With Extended Staffing Variable</u>				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Securities regulator staffing level (extended sample)					0.20 ^b (0.09)	0.25 ^b (0.10)	0.21 ^b (0.09)	0.25 ^a (0.08)	0.24 ^a (0.08)
Formal Public Enforcement	4.92 (3.15)	4.11 (2.80)	5.62 (3.58)	4.64 (3.12)		-4.57 (3.01)			
Disclosure		2.38 (2.85)		3.76 (3.38)			-2.12 (3.90)		0.42 (4.18)
Liability Standards			-1.97 (2.14)	-2.79 (2.55)				-4.47 ^b (2.06)	-4.58 ^c (2.49)
Anti-Directors Rights (rev.)	0.85 (0.79)	0.73 (0.75)	0.97 (0.86)	0.83 (0.80)	0.03 (0.69)	0.22 (0.56)	0.10 (0.58)	0.23 (0.72)	0.22 (0.67)
Judicial Efficiency	0.49 (0.30)	0.42 (0.34)	0.48 (0.31)	0.36 (0.37)	0.31 (0.31)	0.17 (0.27)	0.41 (0.46)	0.29 (0.33)	0.27 (0.50)
Wealth Control	1.01 (0.70)	1.04 (0.72)	1.15 (0.76)	1.26 (0.83)	-0.64 (0.57)	-0.96 (0.80)	-0.87 (0.85)	-0.84 (0.69)	-0.80 (0.83)
Observations	36	36	36	36	25	25	25	25	25
MSE	3.72	3.76	3.75	3.77	3.05	2.98	3.10	2.89	2.97
Adjusted R Squared	0.23	0.21	0.21	0.20	0.41	0.44	0.39	0.47	0.44

a = significant at 1%; b = significant at 5 %; and c = significant at 10 %.

Table 6. Budget-based public enforcement in horse race with private enforcement.

In this table, we again present horse races between resource-based public enforcement and private enforcement variables, in this table using the budget-based (extended sample) to measure the intensity of public resource-based enforcement. Panel A uses market capitalization/GDP as the outcome variable, Panel B trading volume/GDP as the outcome variable, Panel C the number of domestic firms/population, and Panel D IPOs/GDP. We run 5 models with each of the four outcome variables. Columns 1 and 2 run the public enforcement variables, with resource-based enforcement alone first in column 1, and then with formal public enforcement, as measured by the securities regulators' powers in column 2. Columns 3, 4, and 5 check for robustness of the resource-based in horse races between public and private enforcement: i.e., the budget-based resource variable runs against disclosure in column 3, then against liability in column 4, and then against both private enforcement measures in column 5. We add three conditioning variables to each model: wealth, judicial efficiency and anti-directors rights. The resource-based measure is significant in all twenty models, notably including those in column 5 with all of the private-oriented enforcement variables. Disclosure is often significant when run without resource-based enforcement, as seen in Table 5, but is significant only twice (in 8 trials) when run against the budget-based indicator of public enforcement intensity, as seen in column 3 and 5 of the four panels in this table. Liability standards turn negative in the face of the resource-based measure of public enforcement in 4 of the 8 models. Robust standard errors are in parentheses; the constant is not reported.

	Panel A <i>Dependent variable = Stock market capitalization/GDP</i>					Panel B <i>Dependent variable = Trading volume/GDP</i>				
	(1)	(2)	(3)	(4)	(5)	(1)	(2)	(3)	(4)	(5)
Regulatory budget (extended sample)	0.80 ^a (0.27)	0.75 ^b (0.26)	0.77 ^b (0.28)	0.80 ^b (0.28)	0.77 ^b (0.29)	0.36 ^a (0.10)	0.30 ^a (0.09)	0.30 ^b (0.11)	0.35 ^a (0.11)	0.30 ^b (0.11)
Formal public enforcement		54.88 (53.14)					70.56 ^c (39.43)			
Disclosure			28.93 (31.66)		30.75 (39.13)			78.08 ^b (32.87)		78.36 ^b (29.77)
Liability Standards				2.56 (27.48)	-4.23 (31.33)				16.64 (30.37)	-0.66 (26.07)
Anti-director rights (rev.)	18.94 (13.11)	16.31 (12.39)	17.23 (13.26)	18.72 (14.72)	17.49 (14.56)	23.10 ^a (7.22)	19.71 ^b (7.21)	18.47 ^b (7.09)	21.66 ^b (8.29)	18.51 ^b (7.74)
Judicial efficiency	-6.88 (7.07)	-4.18 (6.43)	-6.78 (7.24)	-6.86 (7.26)	-6.81 (7.46)	-4.20 (4.49)	-0.72 (5.04)	-3.93 (4.50)	-4.08 (4.69)	-3.93 (4.62)
Wealth control	64.45 ^b (24.04)	61.78 ^b (22.08)	62.91 ^b (24.77)	64.28 ^b (25.07)	63.09 ^b (25.62)	46.79 ^a (10.40)	43.37 ^a (9.16)	42.64 ^a (10.96)	45.71 ^a (10.78)	42.66 ^a (11.37)
Observations	30	30	30	30	30	30	30	30	30	30
MSE	60.76	60.79	61.74	62.01	63.06	41.81	39.64	39.63	42.41	40.48
Adjusted R Squared	0.60	0.60	0.59	0.59	0.57	0.56	0.60	0.60	0.55	0.59

	Panel C <i>Dependent variable = Domestic firms/population</i>					Panel D <i>Dependent variable = IPOs/GDP</i>				
	(1)	(2)	(3)	(4)	(5)	(1)	(2)	(3)	(4)	(5)
Regulatory budget (extended sample)	0.0049 ^b (0.002)	0.0048 ^b (0.002)	0.0042 ^b (0.002)	0.0046 ^b (0.002)	0.0041 ^b (0.002)	0.038 ^a (0.009)	0.044 ^a (0.011)	0.038 ^a (0.010)	0.041 ^a (0.008)	0.041 ^a (0.008)
Formal public enforcement		0.12 (0.79)					-4.13 (3.05)			
Disclosure			0.93 (0.88)		0.79 (0.96)			-0.68 (4.00)		1.57 (4.95)
Liability Standards				0.52 (0.58)	0.34 (0.62)				-3.08 ^c (1.48)	-3.54 (2.29)
Anti-director rights (rev.)	0.25 (0.19)	0.24 (0.21)	0.19 (0.21)	0.20 (0.20)	0.17 (0.22)	0.17 (0.89)	0.32 (0.77)	0.19 (0.80)	0.30 (0.89)	0.27 (0.86)
Judicial efficiency	0.12 (0.08)	0.13 (0.11)	0.12 (0.08)	0.12 (0.09)	0.13 (0.09)	-0.01 (0.40)	-0.15 (0.36)	0.02 (0.55)	0.01 (0.40)	-0.06 (0.59)
Wealth control	0.83 ^a (0.22)	0.82 ^a (0.23)	0.78 ^a (0.24)	0.79 ^a (0.22)	0.76 ^a (0.23)	1.20 (0.78)	1.17 (0.81)	1.16 (0.95)	1.11 (0.71)	1.19 (0.94)
Observations	30	30	30	30	30	22	22	22	22	22
MSE	0.83	0.85	0.83	0.84	0.84	2.95	2.88	3.04	2.90	2.98
Adjusted R Squared	0.61	0.59	0.61	0.61	0.60	0.51	0.53	0.48	0.52	0.50

a = significant at 1%; b = significant at 5%; and c = significant at 10 %.

Table 7. Distribution of coefficients on public enforcement and private enforcement variables in predicting four financial market size outcomes, when winsorizing and using robust regressions.

In this table, we summarize the coefficients on the resource-based public and the private enforcement variables, with a view to testing for their robustness when influential observations are dampened by winsorizing and by using STATA's rreg command. Columns 1-3, in Panel A, summarize the nature of the coefficients in 24 full horse-race regressions that contain both resource-based and private enforcement variables. We use six resource-based samples, as outlined in Table 1 and whose data is listed in Table 2: We use three samples for each of a staff-based resource measure and a budget-based resource measure. For each of these two we have 3 sub-samples---basic, extended, and extrapolated --- as described in Table 2. These resource-based variables predict four basic outcomes: stock market capitalization, trading volume, number of firms, and IPOs, yielding 24 regressions for the 6 samples and 4 outcomes. Columns 1-3 in Panel A report the overall results for the 3 types of samples and column 4 summarizes those results: The resource-based measures of public enforcement are significant 88% of the time and always positive. The private enforcement variable for disclosure is largely significant (in 42% of the regressions) and always positive. The liability index does less well: never positive and significant and negative more often than positive. In Panel B, columns 5-8 summarize the effect of winsorizing (at the .1 level) the dependent variables and the resource-based enforcement measures. Public enforcement is significant in somewhat more than half of the 24 runs; disclosure is significant in somewhat less than half, and liability is never significant and positive. In Panel C, we report the outcomes that running STATA's rreg command (to clip or eliminate influential observations) has on the results. The resource-based and the disclosure variables hold up fairly well --- significant in nearly half of the runs; the liability index does not, as it's again never significant and positive.

	Panel A Standard OLS regressions				Panel B Winsorized regressions				Panel C Rreg regressions			
	Columns (1)-(4) report the nature of the coefficients in 24 OLS regressions, with 4 financial market size variables as the outcomes and the six resource-based public enforcement variables as the main independent variables				Columns (5)-(8) report the nature of the coefficients in 24 OLS regressions, with 4 winsorized (at the .1 level) financial market size variables as the outcomes and the six similarly winsorized resource-based public enforcement variables as the main independent variables				Columns (9)-(12) report the nature of the coefficients in 24 rreg regressions, with 4 financial market size variables as the outcomes and the six resource-based public enforcement variables as the main independent variables			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
	<i>Coefficients for budget and staffing (basic)</i>	<i>Coefficients for budget and staffing (extended)</i>	<i>Coefficients for budget and staffing (extrapolated)</i>	<i>Subtotal of coefficients in all standard OLS regressions</i>	<i>Coefficients for budget and staffing (basic)</i>	<i>Coefficients for budget and staffing (extended)</i>	<i>Coefficients for budget and staffing (extrapolated)</i>	<i>Subtotal of coefficients in all winsorized regressions</i>	<i>Coefficients for budget and staffing (basic)</i>	<i>Coefficients for budget and staffing (extended)</i>	<i>Coefficients for budget and staffing (extrapolated)</i>	<i>Subtotal of coefficients in all rreg regressions</i>
Resource-based public enforcement												
Coefficients that are both positive and significant	100%	88%	75%	88%	63%	50%	50%	54%	25%	38%	75%	46%
Positive coefficients	100%	100%	100%	100%	100%	100%	88%	96%	100%	88%	88%	92%
Negative coefficients	0%	0%	0%	0%	0%	0%	13%	4%	0%	0%	13%	8%
Coefficients that are both negative and significant	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%
Disclosure index												
Coefficients that are both positive and significant	63%	25%	38%	42%	50%	38%	50%	46%	38%	38%	63%	46%
Positive coefficients	100%	100%	100%	100%	100%	100%	100%	100%	100%	100%	100%	100%
Negative coefficients	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%
Coefficients that are both negative and significant	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%
Liability index												
Coefficients that are both positive and significant	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%
Positive coefficients	25%	25%	50%	33%	38%	63%	75%	58%	50%	50%	75%	58%
Negative coefficients	75%	75%	50%	67%	63%	38%	25%	42%	50%	50%	25%	42%
Coefficients that are both negative and significant	38%	13%	13%	21%	38%	0%	13%	17%	13%	13%	25%	17%
Anti-director rights index (revised)												
Coefficients that are both positive and significant	50%	13%	13%	25%	25%	0%	0%	8%	38%	63%	50%	50%
Positive coefficients	100%	100%	100%	100%	88%	75%	75%	79%	100%	88%	100%	96%
Negative coefficients	0%	0%	0%	0%	0%	25%	25%	21%	0%	13%	0%	4%
Coefficients that are both negative and significant	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%	0%
Judicial efficiency index												
Coefficients that are both positive and significant	13%	13%	25%	17%	13%	13%	25%	17%	25%	13%	25%	21%
Positive coefficients	25%	50%	75%	50%	38%	88%	100%	75%	50%	63%	63%	58%
Negative coefficients	75%	50%	25%	50%	63%	13%	0%	25%	50%	38%	38%	42%
Coefficients that are both negative and significant	13%	0%	0%	4%	13%	0%	0%	4%	0%	0%	25%	8%

Table 8. Regressions with dependent variables associated with private control.

In this table, we examine resource-based public enforcement and private enforcement measures' association with variables related to private control. (These variables were outlined and highlighted in LLS (2006)). Columns 1 and 2 look at block premiums, a variable that seeks to measure the degree to which insiders can transfer value to themselves. Columns 3 and 4 measure access to equity. The outcome measured in 5 and 6 is ownership concentration. Columns 1, 3, and 5 use staffing (extended sample) as the resource-based measure of public enforcement. Columns 2, 4, and 6 use budget (extended sample). Disclosure is significant across all six and in the expected direction. Public enforcement, via the liability index, does not predict these outcomes. Robust standard errors are in parentheses; the constant and the wealth control are not reported.

	<u>Block Premia</u>		<u>Access to Equity</u>		<u>Ownership Concentration</u>	
	(1)	(2)	(3)	(4)	(5)	(6)
	Staff (extended) as resource- based measure	Budget (extended) as resource-based measure	Staff (extended) as resource- based measure	Budget (extended) as resource-based measure	Staff (extended) as resource-based measure	Budget (extended) as resource-based measure
Resource-based public enforcement	0.0012 (0.0014)	0.0003 (0.0002)	-0.0049 (0.0114)	-0.0001 (0.0018)	0.004 ^b (0.0018)	0.0004 (0.0004)
Disclosure	-0.38 ^b (0.14)	-0.38 ^b (0.15)	1.42 ^a (0.47)	1.01 ^b (0.47)	-0.21 ^b (0.10)	-0.31 ^a (0.09)
Liability standards	-0.04 (0.07)	-0.05 (0.07)	0.82 (0.52)	0.85 (0.51)	-0.07 (0.07)	-0.03 (0.06)
Anti-director rights (rev.)	0.02 (0.04)	0.01 (0.04)	-0.17 ^c (0.09)	-0.17 (0.10)	-0.01 (0.01)	-0.01 (0.02)
Judicial efficiency	-0.01 (0.01)	0.00 (0.01)	0.15 ^c (0.07)	0.12 (0.11)	-0.03 ^b (0.01)	-0.03 ^b (0.01)
Observations	25	23	31	27	34	30
MSE	0.10	0.10	0.70	0.71	0.10	0.10
Adjusted R ²	0.30	0.32	0.51	0.47	0.42	0.44

a = significant at 1%; b = significant at 5 %; and c = significant at 10 %.

Table 9. Regressions with new World Bank indices as dependent variables.

In this table we use the World Bank's indices on equity market quality and efficiency as the outcome variables. We regress the outcomes on resource-based public enforcement (via the extended staffing sample in columns 1, 3, 5, 7, and 9, and via the extended budget sample in columns 2, 4, 6, 8, and 10) and on the disclosure and liability private enforcement measures. The resource-based samples are significant in the expected direction in 4 of the 10 runs; disclosure is significant in the expected direction in 6 of 10. But neither has the expected sign (in one instance for each, significantly) in most remaining runs. Liability does less well. Robust standard errors are in parentheses; the constant and the wealth control are not reported.

	<u>Equity Index</u>		<u>Equity Market Size</u>		<u>Market Access</u>		<u>Market Efficiency</u>		<u>Market Stability</u>	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Staff (extended) as resource- based measure	Budget (extended) as resource- based measure	Staff (extended) as resource- based measure	Budget (extended) as resource- based measure	Staff (extended) as resource- based measure	Budget (extended) as resource- based measure	Staff (extended) as resource- based measure	Budget (extended) as resource- based measure	Staff (extended) as resource- based measure	Budget (extended) as resource- based measure
Resource-based public enforcement	0.025 ^c (0.014)	0.003 ^c (0.002)	0.11 ^c (0.059)	0.02 ^a (0.006)	-0.0023 (0.015)	-0.0008 (0.002)	-0.0075 (0.016)	-0.0044 ^c (0.002)	-0.0029 (0.008)	0.0010 (0.001)
Disclosure	0.89 ^b (0.40)	0.84 ^c (0.49)	3.15 ^a (0.96)	2.94 ^b (1.14)	2.02 ^b (0.78)	2.32 ^b (0.85)	-2.21 ^c (1.01)	-1.54 (1.11)	0.47 (0.73)	-0.37 (0.51)
Liability standards	0.09 (0.49)	0.35 (0.43)	-1.31 (1.31)	-0.42 (1.02)	0.36 (0.74)	0.41 (0.73)	0.69 (0.75)	0.69 (0.71)	0.73 ^c (0.40)	0.71 ^c (0.36)
Anti-directors rights (rev.)	0.12 (0.10)	0.20 ^c (0.11)	0.30 (0.32)	0.61 ^c (0.34)	0.04 (0.17)	0.07 (0.18)	0.27 (0.23)	0.20 (0.17)	-0.04 (0.09)	-0.08 (0.08)
Judicial efficiency	0.04 (0.06)	-0.01 (0.07)	-0.04 (0.16)	-0.26 (0.21)	0.09 (0.10)	0.08 (0.11)	0.06 (0.10)	0.11 (0.13)	0.06 (0.05)	0.04 (0.05)
Observations	34	30	34	30	34	30	33	30	34	30
MSE	0.63	0.65	1.76	1.73	1.08	1.11	1.06	1.03	0.53	0.43
Adjusted R ²	0.62	0.62	0.56	0.61	0.11	0.08	0.43	0.51	0.28	0.27

a = significant at 1%; b = significant at 5%; and c = significant at 10%.

Table 10. Public enforcement and common law origins.

This table regresses the resource-based enforcement variables, the formal public enforcement variable, and the private enforcement variables on a common law dummy. The disclosure and liability variables are significant, as prior work has indicated. But the common law dummy is significant as well for both the resource-based public enforcement and the formal powers enforcement variables in half of the measures, although these are not measures traditionally associated with common law institutions. The wealth control here is the World Bank 2005 index measuring the log of gross national income per capita adjusted for purchasing power parity, and log GDP for 2004, which we include to control for scale effects. Robust standard errors are in parentheses; the constant is not reported.

	Panel A <i>Resource-based public enforcement's association with common law</i>						Panel B <i>Private enforcement measures' association with common law</i>			Panel C <i>Public institutional measures' association with common law</i>	
	<i>Basic samples</i>		<i>Extended samples</i>		<i>Extrapolated samples</i>		Disclosure	Liability	Anti-director rights (revised)	Judicial efficiency	Formal public enforcement powers (LLS)
	Securities regulator's staff/population (basic)	Securities regulator's budget/GDP (basic)	Securities regulator's staff/population (extended)	Securities regulator's budget/GDP (extended)	Securities regulator's staff/population (extrapolated)	Securities regulator's budget/GDP (extrapolated)					
Common law dummy	10.88 ^b (4.55)	64.39 (46.04)	6.86 (8.32)	54.13 (38.33)	12.80 ^c (6.40)	60.80 ^c (30.62)	0.32 ^a (0.05)	0.21 ^a (0.06)	1.68 ^a (0.27)	1.27 ^a (0.44)	0.16 ^b (0.07)
Log GNI per capita (purchasing power parity, 2005)	9.54 ^a (3.01)	5.97 (38.43)	27.39 (16.93)	24.45 (37.43)	25.10 ^c (13.19)	17.67 (32.56)	0.03 (0.03)	0.05 (0.03)	0.10 (0.15)	1.62 ^a (0.31)	-0.02 (0.04)
Log GDP (2004)	-2.41 (1.50)	-21.46 (19.87)	-13.39 (10.14)	-33.21 ^c (19.38)	-12.81 (8.17)	-28.90 (17.34)	0.050 ^b (0.02)	0.03 (0.02)	0.15 (0.10)	-0.28 (0.17)	0.01 (0.03)
Observations	30	26	39	35	53	46	49	49	49	49	49
MSE	9.10	92.77	44.29	104.94	38.96	92.84	0.17	0.23	1.04	1.51	0.21
Adjusted R-squared	0.47	0.12	0.21	0.15	0.22	0.15	0.49	0.17	0.36	0.46	0.08

a = significant at 1%; b = significant at 5 %; and c = significant at 10 %.

Table 11. Distribution of coefficients on common law dummy, after controlling for public and private enforcement measures.

This table reports the residual value of the common law dummy, after one controls for public enforcement, as measured with the values reported in Table 2, and private enforcement, using the measures indicated in Table 1. The outcomes are again the same market size measures: stock market capitalization, trading volume, number of domestic firms, and IPOs. This set of specifications yields 24 models. Columns (1)-(4) report the standard OLS regressions, with the four financial market breadth variables as the outcomes. Column 1 reports the distribution of coefficients for the common law dummy, using the basic samples for staffing and for budget as the public enforcement measures. Column 2 does the same, but uses the extended resource-based measures. Column 3 does the same for the extrapolated samples. The last column summarizes the overall results. Overall, the residual effect of common law is never positive and significant, is often negative, and is occasionally negative and significant. The resource-based public enforcement variables are significant in nearly all of the regressions; disclosure is significant in half; liability is never positive and significant.

	Columns (1)-(4) report the nature of the coefficients in 24 OLS regressions, with 4 financial public size variables as the outcomes and the six resource-based public enforcement variables as the main independent variables			
	(1) <i>Coefficients for budget and staffing (basic)</i>	(2) <i>Coefficients for budget and staffing (extended)</i>	(3) <i>Coefficients for budget and staffing (extrapolated)</i>	(4) <i>Subtotal of coefficients in all standard OLS regressions</i>
Resource-based public enforcement				
Coefficients that are both positive and significant	100%	100%	75%	92%
Positive coefficients	100%	100%	100%	100%
Negative coefficients	0%	0%	0%	0%
Coefficients that are both negative and significant	0%	0%	0%	0%
Common law dummy				
Coefficients that are both positive and significant	0%	0%	0%	0%
Positive coefficients	25%	13%	25%	21%
Negative coefficients	75%	88%	75%	79%
Coefficients that are both negative and significant	25%	13%	0%	13%
Disclosure index				
Coefficients that are both positive and significant	75%	38%	38%	50%
Positive coefficients	88%	100%	100%	96%
Negative coefficients	13%	0%	0%	4%
Coefficients that are both negative and significant	0%	0%	0%	0%
Liability index				
Coefficients that are both positive and significant	0%	0%	0%	0%
Positive coefficients	25%	13%	50%	29%
Negative coefficients	75%	88%	50%	71%
Coefficients that are both negative and significant	38%	25%	13%	25%
Anti-director rights index (revised)				
Coefficients that are both positive and significant	63%	38%	13%	38%
Positive coefficients	100%	100%	100%	100%
Negative coefficients	0%	0%	0%	0%
Coefficients that are both negative and significant	0%	0%	0%	0%
Judicial efficiency index				
Coefficients that are both positive and significant	13%	25%	25%	21%
Positive coefficients	50%	75%	88%	71%
Negative coefficients	50%	25%	13%	29%
Coefficients that are both negative and significant	0%	0%	0%	0%

Table 12. Instrumental variable regressions: size of economy (via log GDP).

In this table we use the size of a nation's economy as an instrument for resource-based enforcement in predicting four basic financial outcomes. Panel B reports the first-stage regressions; Panel A reports the second-stage regressions. Panel A-1 examines the 2-stage results with stock market capitalization/GDP as the outcome; Panel A-2 examines trading volume/GDP; Panel A-3 examines the number of domestic firms/population; and Panel A-4 examines IPOs/GDP. Each reports OLS and 2-stage results, with the first set of columns reporting the OLS and 2-stage results when using regulatory staff/population (extended sample). The second set reports the OLS and 2-stage results when using regulatory budget/GDP (extended sample). Robust standard errors are in parentheses; constants and coefficients for the controls as in column 9 of Table 5, are not reported.

<i>Panel A: Second-stage Regressions</i>																
	Panel A-1: Dependent variable = Stock market capitalization/GDP				Panel A-2: Dependent variable = Trading volume/GDP				Panel A-3: Dependent variable = Domestic firms/population				Panel A-4: Dependent variable = IPOs/GDP			
	Instrumented variable is regulatory staff/population (extended sample)		Instrumented variable is regulatory budget/GDP (extended sample)		Instrumented variable is regulatory staff/population (extended sample)		Instrumented variable is regulatory budget/GDP (extended sample)		Instrumented variable is regulatory staff/population (extended sample)		Instrumented variable is regulatory budget/GDP (extended sample)		Instrumented variable is regulatory staff/population (extended sample)		Instrumented variable is regulatory budget/GDP (extended sample)	
	OLS Regression	Ivreg2	OLS Regression	Ivreg2	OLS Regression	Ivreg2	OLS Regression	Ivreg2	OLS Regression	Ivreg2	OLS Regression	Ivreg2	OLS Regression	Ivreg2	OLS Regression	Ivreg2
Resource-based enforcement	6.23 ^b (2.33)	5.48 ^b (2.20)	0.77 ^b (0.29)	0.93 ^b (0.44)	2.11 (1.26)	-1.24 (1.99)	0.30 ^b (0.11)	-0.24 (0.42)	0.03 ^a (0.01)	0.10 ^b (0.04)	0.0041 ^b (0.00)	0.017 ^c (0.01)	0.24 ^a (0.08)	0.22 ^b (0.11)	0.04 ^a (0.01)	0.039 ^b (0.02)
Observations	34	34	30	30	34	34	30	30	34	34	30	28	25	25	22	22
Cragg-Donald Wald F Statistic	n.a.	7.588	n.a.	2.587	n.a.	7.588	n.a.	2.587	n.a.	7.588	n.a.	2.743	n.a.	6.437	n.a.	6.021

<i>Panel B: First-stage Regressions</i>																
	Extended Staff		Extended Budget		Extended Staff		Extended Budget		Extended Staff		Extended Budget					
	OLS	Ivreg2	OLS	Ivreg2	OLS	Ivreg2	OLS	Ivreg2	OLS	Ivreg2	OLS	Ivreg2				
Log GDP in 2005	-3.60 ^b (1.31)	-3.60 ^b (1.31)	-21.34 (12.65)	-21.34 (12.65)	-3.60 ^b (1.31)	-3.60 ^b (1.31)	-20.34 (12.65)	-20.34 (12.65)	-3.60 ^b (1.31)	-3.60 ^b (1.31)	-20.34 (12.65)	-20.34 (12.65)	-3.94 ^b (1.56)	-3.94 ^b (1.56)	-29.47 ^b (12.01)	-29.47 ^b (12.01)
Uncentered R Squared	0.79	0.79	0.71	0.71	0.79	0.79	0.71	0.71	0.79	0.79	0.71	0.71	0.82	0.82	0.73	0.73
F-Test for Excluded Instruments	7.59	7.59	2.59	2.59	7.59	7.59	2.59	2.59	7.59	7.59	2.59	2.59	6.34	6.34	6.02	6.02

a = significant at 1%; b = significant at 5%; and c = significant at 10 %.
 * = Exceeds Stock Yogo Weak ID Test critical values at 15% maximal IV size.